

SOCIETY OF ACTUARIES/CASUALTY ACTUARIAL SOCIETY

EXAM P/1 PROBABILITY

P/1 SAMPLE EXAM SOLUTIONS

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Some of the questions in this study note are taken from past SOA/CAS examinations.

1. Solution: D

Let

G = event that a viewer watched gymnastics

B = event that a viewer watched baseball

S = event that a viewer watched soccer

Then we want to find

$$\begin{aligned}\Pr[(G \cup B \cup S)^c] &= 1 - \Pr(G \cup B \cup S) \\ &= 1 - [\Pr(G) + \Pr(B) + \Pr(S) - \Pr(G \cap B) - \Pr(G \cap S) - \Pr(B \cap S) + \Pr(G \cap B \cap S)] \\ &= 1 - (0.28 + 0.29 + 0.19 - 0.14 - 0.10 - 0.12 + 0.08) = 1 - 0.48 = 0.52\end{aligned}$$

2. Solution: A

Let R = event of referral to a specialist

L = event of lab work

We want to find

$$\begin{aligned}P[R \cap L] &= P[R] + P[L] - P[R \cup L] = P[R] + P[L] - 1 + P[\sim(R \cup L)] \\ &= P[R] + P[L] - 1 + P[\sim R \cap \sim L] = 0.30 + 0.40 - 1 + 0.35 = 0.05.\end{aligned}$$

3. Solution: D

First note

$$P[A \cup B] = P[A] + P[B] - P[A \cap B]$$

$$P[A \cup B'] = P[A] + P[B'] - P[A \cap B']$$

Then add these two equations to get

$$P[A \cup B] + P[A \cup B'] = 2P[A] + (P[B] + P[B']) - (P[A \cap B] + P[A \cap B'])$$

$$0.7 + 0.9 = 2P[A] + 1 - P[(A \cap B) \cup (A \cap B')]$$

$$1.6 = 2P[A] + 1 - P[A]$$

$$P[A] = 0.6$$

4. Solution: A

For $i = 1, 2$, let

R_i = event that a red ball is drawn from urn i

B_i = event that a blue ball is drawn from urn i .

Then if x is the number of blue balls in urn 2,

$$\begin{aligned} 0.44 &= \Pr[(R_1 \cap R_2) \cup (B_1 \cap B_2)] = \Pr[R_1 \cap R_2] + \Pr[B_1 \cap B_2] \\ &= \Pr[R_1] \Pr[R_2] + \Pr[B_1] \Pr[B_2] \\ &= \frac{4}{10} \left(\frac{16}{x+16} \right) + \frac{6}{10} \left(\frac{x}{x+16} \right) \end{aligned}$$

Therefore,

$$\begin{aligned} 2.2 &= \frac{32}{x+16} + \frac{3x}{x+16} = \frac{3x+32}{x+16} \\ 2.2x + 35.2 &= 3x + 32 \\ 0.8x &= 3.2 \\ x &= 4 \end{aligned}$$

5. Solution: D

Let $N(C)$ denote the number of policyholders in classification C . Then

$$\begin{aligned} N(\text{Young} \cap \text{Female} \cap \text{Single}) &= N(\text{Young} \cap \text{Female}) - N(\text{Young} \cap \text{Female} \cap \text{Married}) \\ &= N(\text{Young}) - N(\text{Young} \cap \text{Male}) - [N(\text{Young} \cap \text{Married}) - N(\text{Young} \cap \text{Married} \cap \text{Male})] \\ &= 3000 - 1320 - (1400 - 600) = 880. \end{aligned}$$

6. Solution: B

Let

H = event that a death is due to heart disease

F = event that at least one parent suffered from heart disease

Then based on the medical records,

$$P[H \cap F^c] = \frac{210 - 102}{937} = \frac{108}{937}$$

$$P[F^c] = \frac{937 - 312}{937} = \frac{625}{937}$$

$$\text{and } P[H | F^c] = \frac{P[H \cap F^c]}{P[F^c]} = \frac{108}{937} \bigg/ \frac{625}{937} = \frac{108}{625} = 0.173$$

7. Solution: D

Let

A = event that a policyholder has an auto policy

H = event that a policyholder has a homeowners policy

Then based on the information given,

$$\Pr(A \cap H) = 0.15$$

$$\Pr(A \cap H^c) = \Pr(A) - \Pr(A \cap H) = 0.65 - 0.15 = 0.50$$

$$\Pr(A^c \cap H) = \Pr(H) - \Pr(A \cap H) = 0.50 - 0.15 = 0.35$$

and the portion of policyholders that will renew at least one policy is given by

$$\begin{aligned} & 0.4 \Pr(A \cap H^c) + 0.6 \Pr(A^c \cap H) + 0.8 \Pr(A \cap H) \\ &= (0.4)(0.5) + (0.6)(0.35) + (0.8)(0.15) = 0.53 \quad (= 53\%) \end{aligned}$$

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8. Solution: D

Let

C = event that patient visits a chiropractor

T = event that patient visits a physical therapist

We are given that

$$\Pr[C] = \Pr[T] + 0.14$$

$$\Pr(C \cap T) = 0.22$$

$$\Pr(C^c \cap T^c) = 0.12$$

Therefore,

$$\begin{aligned} 0.88 &= 1 - \Pr[C^c \cap T^c] = \Pr[C \cup T] = \Pr[C] + \Pr[T] - \Pr[C \cap T] \\ &= \Pr[T] + 0.14 + \Pr[T] - 0.22 \\ &= 2 \Pr[T] - 0.08 \end{aligned}$$

or

$$\Pr[T] = (0.88 + 0.08)/2 = 0.48$$

9. Solution: B

Let

M = event that customer insures more than one car

S = event that customer insures a sports car

Then applying DeMorgan's Law, we may compute the desired probability as follows:

$$\begin{aligned}\Pr(M^c \cap S^c) &= \Pr[(M \cup S)^c] = 1 - \Pr(M \cup S) = 1 - [\Pr(M) + \Pr(S) - \Pr(M \cap S)] \\ &= 1 - \Pr(M) - \Pr(S) + \Pr(S|M)\Pr(M) = 1 - 0.70 - 0.20 + (0.15)(0.70) = 0.205\end{aligned}$$

10. Solution: C

Consider the following events about a randomly selected auto insurance customer:

A = customer insures more than one car

B = customer insures a sports car

We want to find the probability of the complement of A intersecting the complement of B (exactly one car, non-sports). But $P(A^c \cap B^c) = 1 - P(A \cup B)$

And, by the Additive Law, $P(A \cup B) = P(A) + P(B) - P(A \cap B)$.

By the Multiplicative Law, $P(A \cap B) = P(B|A)P(A) = 0.15 * 0.64 = 0.096$

It follows that $P(A \cup B) = 0.64 + 0.20 - 0.096 = 0.744$ and $P(A^c \cap B^c) = 0.744 = 0.256$

11. Solution: B

Let

C = Event that a policyholder buys collision coverage

D = Event that a policyholder buys disability coverage

Then we are given that $P[C] = 2P[D]$ and $P[C \cap D] = 0.15$.

By the independence of C and D , it therefore follows that

$$0.15 = P[C \cap D] = P[C]P[D] = 2P[D]P[D] = 2(P[D])^2$$

$$(P[D])^2 = 0.15/2 = 0.075$$

$$P[D] = \sqrt{0.075} \text{ and } P[C] = 2P[D] = 2\sqrt{0.075}$$

Now the independence of C and D also implies the independence of C^c and D^c . As a result, we see that $P[C^c \cap D^c] = P[C^c]P[D^c] = (1 - P[C])(1 - P[D])$

$$= (1 - 2\sqrt{0.075})(1 - \sqrt{0.075}) = 0.33.$$

12. Solution: E

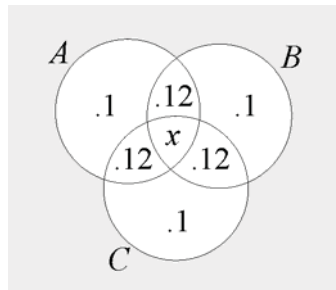
“Boxed” numbers in the table below were computed.

	High BP	Low BP	Norm BP	Total
Regular heartbeat	0.09	0.20	0.56	0.85
Irregular heartbeat	0.05	0.02	0.08	0.15
Total	0.14	0.22	0.64	1.00

From the table, we can see that 20% of patients have a regular heartbeat and low blood pressure.

13. Solution: C

The Venn diagram below summarizes the unconditional probabilities described in the problem.



In addition, we are told that

$$\frac{1}{3} = P[A \cap B \cap C | A \cap B] = \frac{P[A \cap B \cap C]}{P[A \cap B]} = \frac{x}{x + 0.12}$$

It follows that

$$x = \frac{1}{3}(x + 0.12) = \frac{1}{3}x + 0.04$$

$$\frac{2}{3}x = 0.04$$

$$x = 0.06$$

Now we want to find

$$\begin{aligned}
 P[(A \cup B \cup C)^c | A^c] &= \frac{P[(A \cup B \cup C)^c]}{P[A^c]} \\
 &= \frac{1 - P[A \cup B \cup C]}{1 - P[A]} \\
 &= \frac{1 - 3(0.10) - 3(0.12) - 0.06}{1 - 0.10 - 2(0.12) - 0.06} \\
 &= \frac{0.28}{0.60} = 0.467
 \end{aligned}$$

14. Solution: A

$$p_k = \frac{1}{5} p_{k-1} = \frac{1}{5} \frac{1}{5} p_{k-2} = \frac{1}{5} \cdot \frac{1}{5} \cdot \frac{1}{5} p_{k-3} = \dots = \left(\frac{1}{5}\right)^k p_0 \quad k \geq 0$$

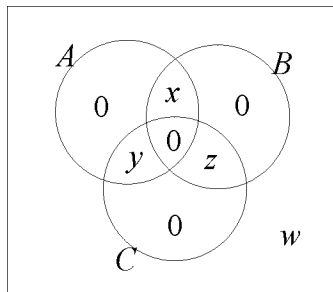
$$1 = \sum_{k=0}^{\infty} p_k = \sum_{k=0}^{\infty} \left(\frac{1}{5}\right)^k p_0 = \frac{p_0}{1 - \frac{1}{5}} = \frac{5}{4} p_0$$

$$p_0 = 4/5 .$$

$$\text{Therefore, } P[N > 1] = 1 - P[N \leq 1] = 1 - (4/5 + 4/5 \cdot 1/5) = 1 - 24/25 = 1/25 = 0.04 .$$

15. Solution: C

A Venn diagram for this situation looks like:



We want to find $w = 1 - (x + y + z)$

$$\text{We have } x + y = \frac{1}{4}, \quad x + z = \frac{1}{3}, \quad y + z = \frac{5}{12}$$

Adding these three equations gives

$$(x + y) + (x + z) + (y + z) = \frac{1}{4} + \frac{1}{3} + \frac{5}{12}$$

$$2(x + y + z) = 1$$

$$x + y + z = \frac{1}{2}$$

$$w = 1 - (x + y + z) = 1 - \frac{1}{2} = \frac{1}{2}$$

Alternatively the three equations can be solved to give $x = 1/12, y = 1/6, z = 1/4$

$$\text{again leading to } w = 1 - \left(\frac{1}{12} + \frac{1}{6} + \frac{1}{4}\right) = \frac{1}{2}$$

16. Solution: D

Let N_1 and N_2 denote the number of claims during weeks one and two, respectively.

Then since N_1 and N_2 are independent,

$$\begin{aligned}\Pr[N_1 + N_2 = 7] &= \sum_{n=0}^7 \Pr[N_1 = n] \Pr[N_2 = 7 - n] \\ &= \sum_{n=0}^7 \left(\frac{1}{2^{n+1}} \right) \left(\frac{1}{2^{8-n}} \right) \\ &= \sum_{n=0}^7 \frac{1}{2^9} \\ &= \frac{8}{2^9} = \frac{1}{2^6} = \frac{1}{64}\end{aligned}$$

17. Solution: D

Let

O = Event of operating room charges

E = Event of emergency room charges

Then

$$\begin{aligned}0.85 &= \Pr(O \cup E) = \Pr(O) + \Pr(E) - \Pr(O \cap E) \\ &= \Pr(O) + \Pr(E) - \Pr(O)\Pr(E) \quad (\text{Independence})\end{aligned}$$

Since $\Pr(E^c) = 0.25 = 1 - \Pr(E)$, it follows $\Pr(E) = 0.75$.

So $0.85 = \Pr(O) + 0.75 - \Pr(O)(0.75)$

$$\Pr(O)(1 - 0.75) = 0.10$$

$$\Pr(O) = 0.40$$

18. Solution: D

Let X_1 and X_2 denote the measurement errors of the less and more accurate instruments, respectively. If $N(\mu, \sigma)$ denotes a normal random variable with mean μ and standard deviation σ , then we are given X_1 is $N(0, 0.0056h)$, X_2 is $N(0, 0.0044h)$ and X_1, X_2 are

independent. It follows that $Y = \frac{X_1 + X_2}{2}$ is $N\left(0, \sqrt{\frac{0.0056^2 h^2 + 0.0044^2 h^2}{4}}\right) = N(0,$

$0.00356h)$. Therefore, $P[-0.005h \leq Y \leq 0.005h] = P[Y \leq 0.005h] - P[Y \leq -0.005h] = P[Y \leq 0.005h] - P[Y \geq 0.005h]$

$$= 2P[Y \leq 0.005h] - 1 = 2P\left[Z \leq \frac{0.005h}{0.00356h}\right] - 1 = 2P[Z \leq 1.4] - 1 = 2(0.9192) - 1 = 0.84.$$

19. Solution: B

Apply Bayes' Formula. Let

A = Event of an accident

B_1 = Event the driver's age is in the range 16-20

B_2 = Event the driver's age is in the range 21-30

B_3 = Event the driver's age is in the range 30-65

B_4 = Event the driver's age is in the range 66-99

Then

$$\begin{aligned}\Pr(B_1|A) &= \frac{\Pr(A|B_1)\Pr(B_1)}{\Pr(A|B_1)\Pr(B_1) + \Pr(A|B_2)\Pr(B_2) + \Pr(A|B_3)\Pr(B_3) + \Pr(A|B_4)\Pr(B_4)} \\ &= \frac{(0.06)(0.08)}{(0.06)(0.08) + (0.03)(0.15) + (0.02)(0.49) + (0.04)(0.28)} = 0.1584\end{aligned}$$

20. Solution: D

Let

S = Event of a standard policy

F = Event of a preferred policy

U = Event of an ultra-preferred policy

D = Event that a policyholder dies

Then

$$\begin{aligned}P[U|D] &= \frac{P[D|U]P[U]}{P[D|S]P[S] + P[D|F]P[F] + P[D|U]P[U]} \\ &= \frac{(0.001)(0.10)}{(0.01)(0.50) + (0.005)(0.40) + (0.001)(0.10)} \\ &= 0.0141\end{aligned}$$

21. Solution: B

Apply Baye's Formula:

$\Pr[\text{Seri.}|\text{Surv.}]$

$$\begin{aligned}&= \frac{\Pr[\text{Surv.}|\text{Seri.}]\Pr[\text{Seri.}]}{\Pr[\text{Surv.}|\text{Crit.}]\Pr[\text{Crit.}] + \Pr[\text{Surv.}|\text{Seri.}]\Pr[\text{Seri.}] + \Pr[\text{Surv.}|\text{Stab.}]\Pr[\text{Stab.}]} \\ &= \frac{(0.9)(0.3)}{(0.6)(0.1) + (0.9)(0.3) + (0.99)(0.6)} = 0.29\end{aligned}$$

22. Solution: D

Let

H = Event of a heavy smoker

L = Event of a light smoker

N = Event of a non-smoker

D = Event of a death within five-year period

Now we are given that $\Pr[D|L] = 2 \Pr[D|N]$ and $\Pr[D|L] = \frac{1}{2} \Pr[D|H]$

Therefore, upon applying Bayes' Formula, we find that

$$\begin{aligned}\Pr[H|D] &= \frac{\Pr[D|H] \Pr[H]}{\Pr[D|N] \Pr[N] + \Pr[D|L] \Pr[L] + \Pr[D|H] \Pr[H]} \\ &= \frac{2 \Pr[D|L](0.2)}{\frac{1}{2} \Pr[D|L](0.5) + \Pr[D|L](0.3) + 2 \Pr[D|L](0.2)} = \frac{0.4}{0.25 + 0.3 + 0.4} = 0.42\end{aligned}$$

23. Solution: D

Let

C = Event of a collision

T = Event of a teen driver

Y = Event of a young adult driver

M = Event of a midlife driver

S = Event of a senior driver

Then using Bayes' Theorem, we see that

$$\begin{aligned}P[Y|C] &= \frac{P[C|Y]P[Y]}{P[C|T]P[T] + P[C|Y]P[Y] + P[C|M]P[M] + P[C|S]P[S]} \\ &= \frac{(0.08)(0.16)}{(0.15)(0.08) + (0.08)(0.16) + (0.04)(0.45) + (0.05)(0.31)} = 0.22.\end{aligned}$$

24. Solution: B

Observe

$$\begin{aligned}\Pr[N \geq 1 | N \leq 4] &= \frac{\Pr[1 \leq N \leq 4]}{\Pr[N \leq 4]} = \frac{\left[\frac{1}{6} + \frac{1}{12} + \frac{1}{20} + \frac{1}{30}\right]}{\left[\frac{1}{2} + \frac{1}{6} + \frac{1}{12} + \frac{1}{20} + \frac{1}{30}\right]} \\ &= \frac{10 + 5 + 3 + 2}{30 + 10 + 5 + 3 + 2} = \frac{20}{50} = \frac{2}{5}\end{aligned}$$

25. Solution: B

Let Y = positive test result

D = disease is present (and $\sim D$ = not D)

Using Baye's theorem:

$$P[D|Y] = \frac{P[Y|D]P[D]}{P[Y|D]P[D] + P[Y|\sim D]P[\sim D]} = \frac{(0.95)(0.01)}{(0.95)(0.01) + (0.005)(0.99)} = 0.657 .$$

26. Solution: C

Let:

S = Event of a smoker

C = Event of a circulation problem

Then we are given that $P[C] = 0.25$ and $P[S|C] = 2 P[S|C^c]$

$$\begin{aligned} \text{Now applying Bayes' Theorem, we find that } P[C|S] &= \frac{P[S|C]P[C]}{P[S|C]P[C] + P[S|C^c]P[C^c]} \\ &= \frac{2P[S|C^c]P[C]}{2P[S|C^c]P[C] + P[S|C^c](1 - P[C])} = \frac{2(0.25)}{2(0.25) + 0.75} = \frac{2}{2+3} = \frac{2}{5} . \end{aligned}$$

27. Solution: D

Use Baye's Theorem with A = the event of an accident in one of the years 1997, 1998 or 1999.

$$\begin{aligned} P[1997|A] &= \frac{P[A|1997]P[1997]}{P[A|1997]P[1997] + P[A|1998]P[1998] + P[A|1999]P[1999]} \\ &= \frac{(0.05)(0.16)}{(0.05)(0.16) + (0.02)(0.18) + (0.03)(0.20)} = 0.45 . \end{aligned}$$

28. Solution: A

Let

C = Event that shipment came from Company X

I_1 = Event that one of the vaccine vials tested is ineffective

Then by Bayes' Formula, $P[C | I_1] = \frac{P[I_1 | C]P[C]}{P[I_1 | C]P[C] + P[I_1 | C^c]P[C^c]}$

Now

$$P[C] = \frac{1}{5}$$

$$P[C^c] = 1 - P[C] = 1 - \frac{1}{5} = \frac{4}{5}$$

$$P[I_1 | C] = \binom{30}{1}(0.10)(0.90)^{29} = 0.141$$

$$P[I_1 | C^c] = \binom{30}{1}(0.02)(0.98)^{29} = 0.334$$

Therefore,

$$P[C | I_1] = \frac{(0.141)(1/5)}{(0.141)(1/5) + (0.334)(4/5)} = 0.096$$

29. Solution: C

Let T denote the number of days that elapse before a high-risk driver is involved in an accident. Then T is exponentially distributed with unknown parameter λ . Now we are given that

$$0.3 = P[T \leq 50] = \int_0^{50} \lambda e^{-\lambda t} dt = -e^{-\lambda t} \Big|_0^{50} = 1 - e^{-50\lambda}$$

Therefore, $e^{-50\lambda} = 0.7$ or $\lambda = -(1/50) \ln(0.7)$

$$\begin{aligned} \text{It follows that } P[T \leq 80] &= \int_0^{80} \lambda e^{-\lambda t} dt = -e^{-\lambda t} \Big|_0^{80} = 1 - e^{-80\lambda} \\ &= 1 - e^{(80/50) \ln(0.7)} = 1 - (0.7)^{80/50} = 0.435. \end{aligned}$$

30. Solution: D

Let N be the number of claims filed. We are given $P[N = 2] = \frac{e^{-\lambda} \lambda^2}{2!} = 3 \frac{e^{-\lambda} \lambda^4}{4!} = 3 \cdot P[N$

$$= 4] 24 \lambda^2 = 6 \lambda^4$$

$$\lambda^2 = 4 \Rightarrow \lambda = 2$$

Therefore, $\text{Var}[N] = \lambda = 2$.

31. Solution: D

Let X denote the number of employees that achieve the high performance level. Then X follows a binomial distribution with parameters $n = 20$ and $p = 0.02$. Now we want to determine x such that

$$\Pr[X > x] \leq 0.01$$

or, equivalently,

$$0.99 \leq \Pr[X \leq x] = \sum_{k=0}^x \binom{20}{k} (0.02)^k (0.98)^{20-k}$$

The following table summarizes the selection process for x :

x	$\Pr[X = x]$	$\Pr[X \leq x]$
0	$(0.98)^{20} = 0.668$	0.668
1	$20(0.02)(0.98)^{19} = 0.272$	0.940
2	$190(0.02)^2(0.98)^{18} = 0.053$	0.993

Consequently, there is less than a 1% chance that more than two employees will achieve the high performance level. We conclude that we should choose the payment amount C such that

$$2C = 120,000$$

or

$$C = 60,000$$

32. Solution: D

Let

X = number of low-risk drivers insured

Y = number of moderate-risk drivers insured

Z = number of high-risk drivers insured

$f(x, y, z)$ = probability function of X , Y , and Z

Then f is a trinomial probability function, so

$$\begin{aligned} \Pr[z \geq x + 2] &= f(0, 0, 4) + f(1, 0, 3) + f(0, 1, 3) + f(0, 2, 2) \\ &= (0.20)^4 + 4(0.50)(0.20)^3 + 4(0.30)(0.20)^3 + \frac{4!}{2!2!}(0.30)^2(0.20)^2 \\ &= 0.0488 \end{aligned}$$

33. Solution: B

Note that

$$\begin{aligned}\Pr[X > x] &= \int_x^{20} 0.005(20-t) dt = 0.005 \left(20t - \frac{1}{2}t^2 \right) \Big|_x^{20} \\ &= 0.005 \left(400 - 200 - 20x + \frac{1}{2}x^2 \right) = 0.005 \left(200 - 20x + \frac{1}{2}x^2 \right)\end{aligned}$$

where $0 < x < 20$. Therefore,

$$\Pr[X > 16 | X > 8] = \frac{\Pr[X > 16]}{\Pr[X > 8]} = \frac{200 - 20(16) + \frac{1}{2}(16)^2}{200 - 20(8) + \frac{1}{2}(8)^2} = \frac{8}{72} = \frac{1}{9}$$

34. Solution: C

We know the density has the form $C(10+x)^{-2}$ for $0 < x < 40$ (equals zero otherwise).

First, determine the proportionality constant C from the condition $\int_0^{40} f(x) dx = 1$:

$$1 = \int_0^{40} C(10+x)^{-2} dx = -C(10+x)^{-1} \Big|_0^{40} = \frac{C}{10} - \frac{C}{50} = \frac{2}{25}C$$

so $C = 25/2$, or 12.5. Then, calculate the probability over the interval (0, 6):

$$12.5 \int_0^6 (10+x)^{-2} dx = -(10+x)^{-1} \Big|_0^6 = \left(\frac{1}{10} - \frac{1}{16} \right) (12.5) = 0.47.$$

35. Solution: C

Let the random variable T be the future lifetime of a 30-year-old. We know that the density of T has the form $f(x) = C(10+x)^{-2}$ for $0 < x < 40$ (and it is equal to zero otherwise). First, determine the proportionality constant C from the condition

$\int_0^{40} f(x) dx = 1$:

$$1 = \int_0^{40} f(x) dx = -C(10+x)^{-1} \Big|_0^{40} = \frac{2}{25}C$$

so that $C = \frac{25}{2} = 12.5$. Then, calculate $P(T < 5)$ by integrating $f(x) = 12.5(10+x)^{-2}$ over the interval (0,5).

36. Solution: B

To determine k, note that

$$1 = \int_0^1 k(1-y)^4 dy = -\frac{k}{5}(1-y)^5 \Big|_0^1 = \frac{k}{5}$$

$$k = 5$$

We next need to find $P[V > 10,000] = P[100,000 Y > 10,000] = P[Y > 0.1]$

$$= \int_{0.1}^1 5(1-y)^4 dy = -(1-y)^5 \Big|_{0.1}^1 = (0.9)^5 = 0.59 \text{ and } P[V > 40,000]$$

$$= P[100,000 Y > 40,000] = P[Y > 0.4] = \int_{0.4}^1 5(1-y)^4 dy = -(1-y)^5 \Big|_{0.4}^1 = (0.6)^5 = 0.078 .$$

It now follows that $P[V > 40,000 | V > 10,000]$

$$= \frac{P[V > 40,000 \cap V > 10,000]}{P[V > 10,000]} = \frac{P[V > 40,000]}{P[V > 10,000]} = \frac{0.078}{0.590} = 0.132 .$$

37. Solution: D

Let T denote printer lifetime. Then $f(t) = \frac{1}{2} e^{-t/2}$, $0 \leq t \leq \infty$

Note that

$$P[T \leq 1] = \int_0^1 \frac{1}{2} e^{-t/2} dt = -e^{-t/2} \Big|_0^1 = 1 - e^{-1/2} = 0.393$$

$$P[1 \leq T \leq 2] = \int_1^2 \frac{1}{2} e^{-t/2} dt = -e^{-t/2} \Big|_1^2 = e^{-1/2} - e^{-1} = 0.239$$

Next, denote refunds for the 100 printers sold by independent and identically distributed random variables Y_1, \dots, Y_{100} where

$$Y_i = \begin{cases} 200 & \text{with probability } 0.393 \\ 100 & \text{with probability } 0.239 \\ 0 & \text{with probability } 0.368 \end{cases} \quad i = 1, \dots, 100$$

$$\text{Now } E[Y_i] = 200(0.393) + 100(0.239) = 102.56$$

$$\text{Therefore, Expected Refunds} = \sum_{i=1}^{100} E[Y_i] = 100(102.56) = 10,256 .$$

38. Solution: A

Let F denote the distribution function of f . Then

$$F(x) = \Pr[X \leq x] = \int_1^x 3t^{-4} dt = -t^{-3} \Big|_1^x = 1 - x^{-3}$$

Using this result, we see

$$\begin{aligned} \Pr[X < 2 | X \geq 1.5] &= \frac{\Pr[(X < 2) \cap (X \geq 1.5)]}{\Pr[X \geq 1.5]} = \frac{\Pr[X < 2] - \Pr[X \leq 1.5]}{\Pr[X \geq 1.5]} \\ &= \frac{F(2) - F(1.5)}{1 - F(1.5)} = \frac{(1.5)^{-3} - (2)^{-3}}{(1.5)^{-3}} = 1 - \left(\frac{3}{4}\right)^3 = 0.578 \end{aligned}$$

39. Solution: E

Let X be the number of hurricanes over the 20-year period. The conditions of the problem give x is a binomial distribution with $n = 20$ and $p = 0.05$. It follows that $P[X < 2] = (0.95)^{20}(0.05)^0 + 20(0.95)^{19}(0.05) + 190(0.95)^{18}(0.05)^2 = 0.358 + 0.377 + 0.189 = 0.925$.

40. Solution: B

Denote the insurance payment by the random variable Y . Then

$$Y = \begin{cases} 0 & \text{if } 0 < X \leq C \\ X - C & \text{if } C < X < 1 \end{cases}$$

Now we are given that

$$0.64 = \Pr(Y < 0.5) = \Pr(0 < X < 0.5 + C) = \int_0^{0.5+C} 2x \, dx = x^2 \Big|_0^{0.5+C} = (0.5 + C)^2$$

Therefore, solving for C , we find $C = \pm 0.8 - 0.5$

Finally, since $0 < C < 1$, we conclude that $C = 0.3$

41. Solution: E

Let

X = number of group 1 participants that complete the study.

Y = number of group 2 participants that complete the study.

Now we are given that X and Y are independent.

Therefore,

$$\begin{aligned}
 & P\{[(X \geq 9) \cap (Y < 9)] \cup [(X < 9) \cap (Y \geq 9)]\} \\
 &= P[(X \geq 9) \cap (Y < 9)] + P[(X < 9) \cap (Y \geq 9)] \\
 &= 2P[(X \geq 9) \cap (Y < 9)] \quad (\text{due to symmetry}) \\
 &= 2P[X \geq 9]P[Y < 9] \\
 &= 2P[X \geq 9]P[X < 9] \quad (\text{again due to symmetry}) \\
 &= 2P[X \geq 9](1 - P[X \geq 9]) \\
 &= 2\left[\binom{10}{9}(0.2)(0.8)^9 + \binom{10}{10}(0.8)^{10}\right]\left[1 - \binom{10}{9}(0.2)(0.8)^9 - \binom{10}{10}(0.8)^{10}\right] \\
 &= 2[0.376][1 - 0.376] = 0.469
 \end{aligned}$$

42. Solution: D

Let

I_A = Event that Company A makes a claim

I_B = Event that Company B makes a claim

X_A = Expense paid to Company A if claims are made

X_B = Expense paid to Company B if claims are made

Then we want to find

$$\begin{aligned}
 & \Pr\{[I_A^c \cap I_B] \cup [(I_A \cap I_B) \cap (X_A < X_B)]\} \\
 &= \Pr[I_A^c \cap I_B] + \Pr[(I_A \cap I_B) \cap (X_A < X_B)] \\
 &= \Pr[I_A^c] \Pr[I_B] + \Pr[I_A] \Pr[I_B] \Pr[X_A < X_B] \quad (\text{independence}) \\
 &= (0.60)(0.30) + (0.40)(0.30) \Pr[X_B - X_A \geq 0] \\
 &= 0.18 + 0.12 \Pr[X_B - X_A \geq 0]
 \end{aligned}$$

Now $X_B - X_A$ is a linear combination of independent normal random variables.

Therefore, $X_B - X_A$ is also a normal random variable with mean

$$M = E[X_B - X_A] = E[X_B] - E[X_A] = 9,000 - 10,000 = -1,000$$

and standard deviation $\sigma = \sqrt{\text{Var}(X_B) + \text{Var}(X_A)} = \sqrt{(2000)^2 + (2000)^2} = 2000\sqrt{2}$

It follows that

$$\begin{aligned}
\Pr[X_B - X_A \geq 0] &= \Pr\left[Z \geq \frac{1000}{2000\sqrt{2}}\right] \quad (Z \text{ is standard normal}) \\
&= \Pr\left[Z \geq \frac{1}{2\sqrt{2}}\right] \\
&= 1 - \Pr\left[Z < \frac{1}{2\sqrt{2}}\right] \\
&= 1 - \Pr[Z < 0.354] \\
&= 1 - 0.638 = 0.362
\end{aligned}$$

Finally,

$$\begin{aligned}
\Pr\left\{\left[I_A^C \cap I_B\right] \cup \left[(I_A \cap I_B) \cap (X_A < X_B)\right]\right\} &= 0.18 + (0.12)(0.362) \\
&= 0.223
\end{aligned}$$

43. Solution: D

If a month with one or more accidents is regarded as success and k = the number of failures before the fourth success, then k follows a negative binomial distribution and the requested probability is

$$\begin{aligned}
\Pr[k \geq 4] &= 1 - \Pr[k \leq 3] = 1 - \sum_{k=0}^3 \binom{3+k}{k} \left(\frac{3}{5}\right)^4 \left(\frac{2}{5}\right)^k \\
&= 1 - \left(\frac{3}{5}\right)^4 \left[\binom{3}{0} \left(\frac{2}{5}\right)^0 + \binom{4}{1} \left(\frac{2}{5}\right)^1 + \binom{5}{2} \left(\frac{2}{5}\right)^2 + \binom{6}{3} \left(\frac{2}{5}\right)^3 \right] \\
&= 1 - \left(\frac{3}{5}\right)^4 \left[1 + \frac{8}{5} + \frac{8}{5} + \frac{32}{25} \right] \\
&= 0.2898
\end{aligned}$$

Alternatively the solution is

$$\left(\frac{2}{5}\right)^4 + \binom{4}{1} \left(\frac{2}{5}\right)^4 \frac{3}{5} + \binom{5}{2} \left(\frac{2}{5}\right)^4 \left(\frac{3}{5}\right)^2 + \binom{6}{3} \left(\frac{2}{5}\right)^4 \left(\frac{3}{5}\right)^3 = 0.2898$$

which can be derived directly or by regarding the problem as a negative binomial distribution with

- i) success taken as a month with no accidents
- ii) k = the number of failures before the fourth success, and
- iii) calculating $\Pr[k \leq 3]$

44. Solution: C

If k is the number of days of hospitalization, then the insurance payment $g(k)$ is

$$g(k) = \begin{cases} 100k & \text{for } k=1, 2, 3 \\ 300+50(k-3) & \text{for } k=4, 5. \end{cases}$$

Thus, the expected payment is $\sum_{k=1}^5 g(k)p_k = 100p_1 + 200p_2 + 300p_3 + 350p_4 + 400p_5 =$

$$\frac{1}{15}(100 \times 5 + 200 \times 4 + 300 \times 3 + 350 \times 2 + 400 \times 1) = 220$$

45. Solution: D

$$\text{Note that } E(X) = \int_{-2}^0 -\frac{x^2}{10} dx + \int_0^4 \frac{x^2}{10} dx = -\frac{x^3}{30} \Big|_{-2}^0 + \frac{x^3}{30} \Big|_0^4 = -\frac{8}{30} + \frac{64}{30} = \frac{56}{30} = \frac{28}{15}$$

46. Solution: D

The density function of T is

$$f(t) = \frac{1}{3}e^{-t/3}, \quad 0 < t < \infty$$

Therefore,

$$\begin{aligned} E[X] &= E[\max(T, 2)] \\ &= \int_0^2 \frac{2}{3}e^{-t/3} dt + \int_2^\infty \frac{t}{3}e^{-t/3} dt \\ &= -2e^{-t/3} \Big|_0^2 - te^{-t/3} \Big|_2^\infty + \int_2^\infty e^{-t/3} dt \\ &= -2e^{-2/3} + 2 + 2e^{-2/3} - 3e^{-t/3} \Big|_2^\infty \\ &= 2 + 3e^{-2/3} \end{aligned}$$

47. Solution: D

Let T be the time from purchase until failure of the equipment. We are given that T is exponentially distributed with parameter $\lambda = 10$ since $10 = E[T] = \lambda$. Next define the payment

$$P \text{ under the insurance contract by } P = \begin{cases} x & \text{for } 0 \leq T \leq 1 \\ \frac{x}{2} & \text{for } 1 < T \leq 3 \\ 0 & \text{for } T > 3 \end{cases}$$

We want to find x such that

$$\begin{aligned} 1000 &= E[P] = \int_0^1 \frac{x}{10} e^{-t/10} dt + \int_1^3 \frac{x}{2} \frac{1}{10} e^{-t/10} dt = -xe^{-t/10} \Big|_0^1 - \frac{x}{2} e^{-t/10} \Big|_1^3 \\ &= -x e^{-1/10} + x - (x/2) e^{-3/10} + (x/2) e^{-1/10} = x(1 - \frac{1}{2} e^{-1/10} - \frac{1}{2} e^{-3/10}) = 0.1772x. \end{aligned}$$

We conclude that $x = 5644$.

48. Solution: E

Let X and Y denote the year the device fails and the benefit amount, respectively. Then the density function of X is given by

$$f(x) = (0.6)^{x-1} (0.4), \quad x = 1, 2, 3, \dots$$

and

$$y = \begin{cases} 1000(5-x) & \text{if } x = 1, 2, 3, 4 \\ 0 & \text{if } x > 4 \end{cases}$$

It follows that

$$\begin{aligned} E[Y] &= 4000(0.4) + 3000(0.6)(0.4) + 2000(0.6)^2(0.4) + 1000(0.6)^3(0.4) \\ &= 2694 \end{aligned}$$

49. Solution: D

Define $f(X)$ to be hospitalization payments made by the insurance policy. Then

$$f(X) = \begin{cases} 100X & \text{if } X = 1, 2, 3 \\ 300 + 25(X-3) & \text{if } X = 4, 5 \end{cases}$$

and

$$\begin{aligned}
E[f(X)] &= \sum_{k=1}^5 f(k) \Pr[X = k] \\
&= 100 \left(\frac{5}{15} \right) + 200 \left(\frac{4}{15} \right) + 300 \left(\frac{3}{15} \right) + 325 \left(\frac{2}{15} \right) + 350 \left(\frac{1}{15} \right) \\
&= \frac{1}{3} [100 + 160 + 180 + 130 + 70] = \frac{640}{3} = 213.33
\end{aligned}$$

50. Solution: C

Let N be the number of major snowstorms per year, and let P be the amount paid to the company under the policy. Then $\Pr[N = n] = \frac{(3/2)^n e^{-3/2}}{n!}$, $n = 0, 1, 2, \dots$ and

$$P = \begin{cases} 0 & \text{for } N = 0 \\ 10,000(N-1) & \text{for } N \geq 1 \end{cases}.$$

$$\begin{aligned}
\text{Now observe that } E[P] &= \sum_{n=1}^{\infty} 10,000(n-1) \frac{(3/2)^n e^{-3/2}}{n!} \\
&= 10,000 e^{-3/2} + \sum_{n=0}^{\infty} 10,000(n-1) \frac{(3/2)^n e^{-3/2}}{n!} = 10,000 e^{-3/2} + E[10,000(N-1)] \\
&= 10,000 e^{-3/2} + E[10,000N] - E[10,000] = 10,000 e^{-3/2} + 10,000(3/2) - 10,000 = 7,231.
\end{aligned}$$

51. Solution: C

Let Y denote the manufacturer's retained annual losses.

$$\text{Then } Y = \begin{cases} x & \text{for } 0.6 < x \leq 2 \\ 2 & \text{for } x > 2 \end{cases}$$

$$\begin{aligned}
\text{and } E[Y] &= \int_{0.6}^2 x \left[\frac{2.5(0.6)^{2.5}}{x^{3.5}} \right] dx + \int_2^{\infty} 2 \left[\frac{2.5(0.6)^{2.5}}{x^{3.5}} \right] dx = \int_{0.6}^2 \frac{2.5(0.6)^{2.5}}{x^{2.5}} dx - \frac{2(0.6)^{2.5}}{x^{2.5}} \Big|_2^{\infty} \\
&= -\frac{2.5(0.6)^{2.5}}{1.5x^{1.5}} \Big|_{0.6}^2 + \frac{2(0.6)^{2.5}}{(2)^{2.5}} = -\frac{2.5(0.6)^{2.5}}{1.5(2)^{1.5}} + \frac{2.5(0.6)^{2.5}}{1.5(0.6)^{1.5}} + \frac{(0.6)^{2.5}}{2^{1.5}} = 0.9343.
\end{aligned}$$

52. Solution: A

Let us first determine K . Observe that

$$1 = K \left(1 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \frac{1}{5} \right) = K \left(\frac{60 + 30 + 20 + 15 + 12}{60} \right) = K \left(\frac{137}{60} \right)$$

$$K = \frac{60}{137}$$

It then follows that

$$\begin{aligned} \Pr[N = n] &= \Pr[N = n | \text{Insured Suffers a Loss}] \Pr[\text{Insured Suffers a Loss}] \\ &= \frac{60}{137N} (0.05) = \frac{3}{137N}, \quad N = 1, \dots, 5 \end{aligned}$$

Now because of the deductible of 2, the net annual premium $P = E[X]$ where

$$X = \begin{cases} 0 & , \text{ if } N \leq 2 \\ N - 2 & , \text{ if } N > 2 \end{cases}$$

Then,

$$P = E[X] = \sum_{N=3}^5 (N-2) \frac{3}{137N} = (1) \left(\frac{1}{137} \right) + 2 \left[\frac{3}{137(4)} \right] + 3 \left[\frac{3}{137(5)} \right] = 0.0314$$

53. Solution: D

Let W denote claim payments. Then $W = \begin{cases} y & \text{for } 1 < y \leq 10 \\ 10 & \text{for } y \geq 10 \end{cases}$

It follows that $E[W] = \int_1^{10} y \frac{2}{y^3} dy + \int_{10}^{\infty} 10 \frac{2}{y^3} dy = -\frac{2}{y} \Big|_1^{10} - \frac{10}{y^2} \Big|_{10}^{\infty} = 2 - 2/10 + 1/10 = 1.9$.

54. Solution: B

Let Y denote the claim payment made by the insurance company.

Then

$$Y = \begin{cases} 0 & \text{with probability } 0.94 \\ \text{Max } (0, x - 1) & \text{with probability } 0.04 \\ 14 & \text{with probability } 0.02 \end{cases}$$

and

$$\begin{aligned} E[Y] &= (0.94)(0) + (0.04)(0.5003) \int_1^{15} (x-1)e^{-x/2} dx + (0.02)(14) \\ &= (0.020012) \left[\int_1^{15} xe^{-x/2} dx - \int_1^{15} e^{-x/2} dx \right] + 0.28 \\ &= 0.28 + (0.020012) \left[-2xe^{-x/2} \Big|_1^{15} + 2 \int_1^{15} e^{-x/2} dx - \int_1^{15} e^{-x/2} dx \right] \\ &= 0.28 + (0.020012) \left[-30e^{-7.5} + 2e^{-0.5} + \int_1^{15} e^{-x/2} dx \right] \\ &= 0.28 + (0.020012) \left[-30e^{-7.5} + 2e^{-0.5} - 2e^{-x/2} \Big|_1^{15} \right] \\ &= 0.28 + (0.020012) (-30e^{-7.5} + 2e^{-0.5} - 2e^{-7.5} + 2e^{-0.5}) \\ &= 0.28 + (0.020012) (-32e^{-7.5} + 4e^{-0.5}) \\ &= 0.28 + (0.020012)(2.408) \\ &= 0.328 \quad (\text{in thousands}) \end{aligned}$$

It follows that the expected claim payment is 328 .

55. Solution: C

The pdf of x is given by $f(x) = \frac{k}{(1+x)^4}$, $0 < x < \infty$. To find k , note

$$1 = \int_0^{\infty} \frac{k}{(1+x)^4} dx = -\frac{k}{3} \frac{1}{(1+x)^3} \Big|_0^{\infty} = \frac{k}{3}$$

$$k = 3$$

It then follows that $E[x] = \int_0^{\infty} \frac{3x}{(1+x)^4} dx$ and substituting $u = 1 + x$, $du = dx$, we see

$$E[x] = \int_1^{\infty} \frac{3(u-1)}{u^4} du = 3 \int_1^{\infty} (u^{-3} - u^{-4}) du = 3 \left[\frac{u^{-2}}{-2} - \frac{u^{-3}}{-3} \right]_1^{\infty} = 3 \left[\frac{1}{2} - \frac{1}{3} \right] = 3/2 - 1 = 1/2 .$$

56. Solution: C

Let Y represent the payment made to the policyholder for a loss subject to a deductible D.

$$\text{That is } Y = \begin{cases} 0 & \text{for } 0 \leq X \leq D \\ x - D & \text{for } D < X \leq 1 \end{cases}$$

Then since $E[X] = 500$, we want to choose D so that

$$\frac{1}{4} 500 = \int_D^{1000} \frac{1}{1000} (x - D) dx = \frac{1}{1000} \left. \frac{(x - D)^2}{2} \right|_D^{1000} = \frac{(1000 - D)^2}{2000}$$

$$(1000 - D)^2 = 2000/4 \cdot 500 = 500^2$$

$$1000 - D = \pm 500$$

$$D = 500 \text{ (or } D = 1500 \text{ which is extraneous).}$$

57. Solution: B

We are given that $M_X(t) = \frac{1}{(1 - 2500t)^4}$ for the claim size X in a certain class of accidents.

$$\text{First, compute } M_X'(t) = \frac{(-4)(-2500)}{(1 - 2500t)^5} = \frac{10,000}{(1 - 2500t)^5}$$

$$M_X''(t) = \frac{(10,000)(-5)(-2500)}{(1 - 2500t)^6} = \frac{125,000,000}{(1 - 2500t)^6}$$

$$\text{Then } E[X] = M_X'(0) = 10,000$$

$$E[X^2] = M_X''(0) = 125,000,000$$

$$\text{Var}[X] = E[X^2] - \{E[X]\}^2 = 125,000,000 - (10,000)^2 = 25,000,000$$

$$\sqrt{\text{Var}[X]} = 5,000.$$

58. Solution: E

Let X_J , X_K , and X_L represent annual losses for cities J, K, and L, respectively. Then $X = X_J + X_K + X_L$ and due to independence

$$M(t) = E[e^{xt}] = E[e^{(x_J + x_K + x_L)t}] = E[e^{x_J t}] E[e^{x_K t}] E[e^{x_L t}]$$

$$= M_J(t) M_K(t) M_L(t) = (1 - 2t)^{-3} (1 - 2t)^{-2.5} (1 - 2t)^{-4.5} = (1 - 2t)^{-10}$$

Therefore,

$$M'(t) = 20(1 - 2t)^{-11}$$

$$M''(t) = 440(1 - 2t)^{-12}$$

$$M'''(t) = 10,560(1 - 2t)^{-13}$$

$$E[X^3] = M'''(0) = 10,560$$

59. Solution: B

The distribution function of X is given by

$$F(x) = \int_{200}^x \frac{2.5(200)^{2.5}}{t^{3.5}} dt = \frac{-(200)^{2.5}}{t^{2.5}} \Big|_{200}^x = 1 - \frac{(200)^{2.5}}{x^{2.5}}, \quad x > 200$$

Therefore, the p^{th} percentile x_p of X is given by

$$\frac{p}{100} = F(x_p) = 1 - \frac{(200)^{2.5}}{x_p^{2.5}}$$

$$1 - 0.01p = \frac{(200)^{2.5}}{x_p^{2.5}}$$

$$(1 - 0.01p)^{2/5} = \frac{200}{x_p}$$

$$x_p = \frac{200}{(1 - 0.01p)^{2/5}}$$

$$\text{It follows that } x_{70} - x_{30} = \frac{200}{(0.30)^{2/5}} - \frac{200}{(0.70)^{2/5}} = 93.06$$

60. Solution: E

Let X and Y denote the annual cost of maintaining and repairing a car before and after the 20% tax, respectively. Then $Y = 1.2X$ and $\text{Var}[Y] = \text{Var}[1.2X] = (1.2)^2 \text{Var}[X] = (1.2)^2(260) = 374$.

61. Solution: A

The first quartile, $Q1$, is found by $\frac{3}{4} = \int_{Q1}^{\infty} f(x) dx$. That is, $\frac{3}{4} = (200/Q1)^{2.5}$ or

$Q1 = 200 (4/3)^{0.4} = 224.4$. Similarly, the third quartile, $Q3$, is given by $Q3 = 200 (4)^{0.4} = 348.2$. The interquartile range is the difference $Q3 - Q1$.

62. Solution: C

First note that the density function of X is given by

$$f(x) = \begin{cases} \frac{1}{2} & \text{if } x = 1 \\ x-1 & \text{if } 1 < x < 2 \\ 0 & \text{otherwise} \end{cases}$$

Then

$$\begin{aligned} E(X) &= \frac{1}{2} + \int_1^2 x(x-1)dx = \frac{1}{2} + \int_1^2 (x^2 - x)dx = \frac{1}{2} + \left(\frac{1}{3}x^3 - \frac{1}{2}x^2 \right) \Big|_1^2 \\ &= \frac{1}{2} + \frac{8}{3} - \frac{4}{2} - \frac{1}{3} + \frac{1}{2} = \frac{7}{3} - 1 = \frac{4}{3} \\ E(X^2) &= \frac{1}{2} + \int_1^2 x^2(x-1)dx = \frac{1}{2} + \int_1^2 (x^3 - x^2)dx = \frac{1}{2} + \left(\frac{1}{4}x^4 - \frac{1}{3}x^3 \right) \Big|_1^2 \\ &= \frac{1}{2} + \frac{16}{4} - \frac{8}{3} - \frac{1}{4} + \frac{1}{3} = \frac{17}{4} - \frac{7}{3} = \frac{23}{12} \\ \text{Var}(X) &= E(X^2) - [E(X)]^2 = \frac{23}{12} - \left(\frac{4}{3} \right)^2 = \frac{23}{12} - \frac{16}{9} = \frac{5}{36} \end{aligned}$$

63. Solution: C

$$\text{Note } Y = \begin{cases} X & \text{if } 0 \leq X \leq 4 \\ 4 & \text{if } 4 < X \leq 5 \end{cases}$$

Therefore,

$$\begin{aligned} E[Y] &= \int_0^4 \frac{1}{5}x dx + \int_4^5 \frac{4}{5}dx = \frac{1}{10}x^2 \Big|_0^4 + \frac{4}{5}x \Big|_4^5 \\ &= \frac{16}{10} + \frac{20}{5} - \frac{16}{5} = \frac{8}{5} + \frac{4}{5} = \frac{12}{5} \\ E[Y^2] &= \int_0^4 \frac{1}{5}x^2 dx + \int_4^5 \frac{16}{5}dx = \frac{1}{15}x^3 \Big|_0^4 + \frac{16}{5}x \Big|_4^5 \\ &= \frac{64}{15} + \frac{80}{5} - \frac{64}{5} = \frac{64}{15} + \frac{16}{5} = \frac{64}{15} + \frac{48}{15} = \frac{112}{15} \\ \text{Var}[Y] &= E[Y^2] - (E[Y])^2 = \frac{112}{15} - \left(\frac{12}{5} \right)^2 = 1.71 \end{aligned}$$

64. Solution: A

Let X denote claim size. Then $E[X] = [20(0.15) + 30(0.10) + 40(0.05) + 50(0.20) + 60(0.10) + 70(0.10) + 80(0.30)] = (3 + 3 + 2 + 10 + 6 + 7 + 24) = 55$
 $E[X^2] = 400(0.15) + 900(0.10) + 1600(0.05) + 2500(0.20) + 3600(0.10) + 4900(0.10) + 6400(0.30) = 60 + 90 + 80 + 500 + 360 + 490 + 1920 = 3500$
 $\text{Var}[X] = E[X^2] - (E[X])^2 = 3500 - 3025 = 475$ and $\sqrt{\text{Var}[X]} = 21.79$.
 Now the range of claims within one standard deviation of the mean is given by $[55.00 - 21.79, 55.00 + 21.79] = [33.21, 76.79]$
 Therefore, the proportion of claims within one standard deviation is $0.05 + 0.20 + 0.10 + 0.10 = 0.45$.

65. Solution: B

Let X and Y denote repair cost and insurance payment, respectively, in the event the auto is damaged. Then

$$Y = \begin{cases} 0 & \text{if } x \leq 250 \\ x - 250 & \text{if } x > 250 \end{cases}$$

and

$$E[Y] = \int_{250}^{1500} \frac{1}{1500} (x - 250) dx = \frac{1}{3000} (x - 250)^2 \Big|_{250}^{1500} = \frac{1250^2}{3000} = 521$$

$$E[Y^2] = \int_{250}^{1500} \frac{1}{1500} (x - 250)^2 dx = \frac{1}{4500} (x - 250)^3 \Big|_{250}^{1500} = \frac{1250^3}{4500} = 434,028$$

$$\text{Var}[Y] = E[Y^2] - \{E[Y]\}^2 = 434,028 - (521)^2$$

$$\sqrt{\text{Var}[Y]} = 403$$

66. Solution: E

Let X_1, X_2, X_3 , and X_4 denote the four independent bids with common distribution function F . Then if we define $Y = \max(X_1, X_2, X_3, X_4)$, the distribution function G of Y is given by

$$\begin{aligned} G(y) &= \Pr[Y \leq y] \\ &= \Pr[(X_1 \leq y) \cap (X_2 \leq y) \cap (X_3 \leq y) \cap (X_4 \leq y)] \\ &= \Pr[X_1 \leq y] \Pr[X_2 \leq y] \Pr[X_3 \leq y] \Pr[X_4 \leq y] \\ &= [F(y)]^4 \\ &= \frac{1}{16} (1 + \sin \pi y)^4, \quad \frac{3}{2} \leq y \leq \frac{5}{2} \end{aligned}$$

It then follows that the density function g of Y is given by

$$\begin{aligned}
g(y) &= G'(y) \\
&= \frac{1}{4}(1 + \sin \pi y)^3 (\pi \cos \pi y) \\
&= \frac{\pi}{4} \cos \pi y (1 + \sin \pi y)^3, \quad \frac{3}{2} \leq y \leq \frac{5}{2}
\end{aligned}$$

Finally,

$$\begin{aligned}
E[Y] &= \int_{3/2}^{5/2} yg(y) dy \\
&= \int_{3/2}^{5/2} \frac{\pi}{4} y \cos \pi y (1 + \sin \pi y)^3 dy
\end{aligned}$$

67. Solution: B

The amount of money the insurance company will have to pay is defined by the random variable

$$Y = \begin{cases} 1000x & \text{if } x < 2 \\ 2000 & \text{if } x \geq 2 \end{cases}$$

where x is a Poisson random variable with mean 0.6. The probability function for X is

$$p(x) = \frac{e^{-0.6} (0.6)^k}{k!} \quad k = 0, 1, 2, 3, \dots \text{ and}$$

$$\begin{aligned}
E[Y] &= 0 + 1000(0.6)e^{-0.6} + 2000e^{-0.6} \sum_{k=2}^{\infty} \frac{0.6^k}{k!} \\
&= 1000(0.6)e^{-0.6} + 2000 \left(e^{-0.6} \sum_{k=0}^{\infty} \frac{0.6^k}{k!} - e^{-0.6} - (0.6)e^{-0.6} \right) \\
&= 2000e^{-0.6} \sum_{k=0}^{\infty} \frac{(0.6)^k}{k!} - 2000e^{-0.6} - 1000(0.6)e^{-0.6} = 2000 - 2000e^{-0.6} - 600e^{-0.6} \\
&= 573 \\
E[Y^2] &= (1000)^2 (0.6)e^{-0.6} + (2000)^2 e^{-0.6} \sum_{k=2}^{\infty} \frac{0.6^k}{k!} \\
&= (2000)^2 e^{-0.6} \sum_{k=0}^{\infty} \frac{0.6^k}{k!} - (2000)^2 e^{-0.6} - [(2000)^2 - (1000)^2] (0.6)e^{-0.6} \\
&= (2000)^2 - (2000)^2 e^{-0.6} - [(2000)^2 - (1000)^2] (0.6)e^{-0.6} \\
&= 816,893
\end{aligned}$$

$$\text{Var}[Y] = E[Y^2] - \{E[Y]\}^2 = 816,893 - (573)^2 = 488,564$$

$$\sqrt{\text{Var}[Y]} = 699$$

68. Solution: C

Note that X has an exponential distribution. Therefore, $c = 0.004$. Now let Y denote the claim benefits paid. Then $Y = \begin{cases} x & \text{for } x < 250 \\ 250 & \text{for } x \geq 250 \end{cases}$ and we want to find m such that 0.50

$$= \int_0^m 0.004e^{-0.004x} dx = -e^{-0.004x} \Big|_0^m = 1 - e^{-0.004m}$$

This condition implies $e^{-0.004m} = 0.5 \Rightarrow m = 250 \ln 2 = 173.29$.

69. Solution: D

The distribution function of an exponential random variable

T with parameter θ is given by $F(t) = 1 - e^{-t/\theta}$, $t > 0$

Since we are told that T has a median of four hours, we may determine θ as follows:

$$\frac{1}{2} = F(4) = 1 - e^{-4/\theta}$$

$$\frac{1}{2} = e^{-4/\theta}$$

$$-\ln(2) = -\frac{4}{\theta}$$

$$\theta = \frac{4}{\ln(2)}$$

Therefore, $\Pr(T \geq 5) = 1 - F(5) = e^{-5/\theta} = e^{-\frac{5\ln(2)}{4}} = 2^{-5/4} = 0.42$

70. Solution: E

Let X denote actual losses incurred. We are given that X follows an exponential distribution with mean 300, and we are asked to find the 95th percentile of all claims that exceed 100. Consequently, we want to find p_{95} such that

$$0.95 = \frac{\Pr[100 < x < p_{95}]}{P[X > 100]} = \frac{F(p_{95}) - F(100)}{1 - F(100)} \text{ where } F(x) \text{ is the distribution function of } X.$$

Now $F(x) = 1 - e^{-x/300}$.

$$\text{Therefore, } 0.95 = \frac{1 - e^{-p_{95}/300} - (1 - e^{-100/300})}{1 - (1 - e^{-100/300})} = \frac{e^{-1/3} - e^{-p_{95}/300}}{e^{-1/3}} = 1 - e^{1/3} e^{-p_{95}/300}$$

$$e^{-p_{95}/300} = 0.05 e^{-1/3}$$

$$p_{95} = -300 \ln(0.05 e^{-1/3}) = 999$$

71. Solution: A

The distribution function of Y is given by

$$G(y) = \Pr(T^2 \leq y) = \Pr(T \leq \sqrt{y}) = F(\sqrt{y}) = 1 - 4/y$$

for $y > 4$. Differentiate to obtain the density function $g(y) = 4y^{-2}$

Alternate solution:

Differentiate $F(t)$ to obtain $f(t) = 8t^{-3}$ and set $y = t^2$. Then $t = \sqrt{y}$ and

$$g(y) = f(t(y))|dt/dy| = f(\sqrt{y})\left|\frac{d}{dt}(\sqrt{y})\right| = 8y^{-3/2}\left(\frac{1}{2}y^{-1/2}\right) = 4y^{-2}$$

72. Solution: E

We are given that R is uniform on the interval $(0.04, 0.08)$ and $V = 10,000e^R$

Therefore, the distribution function of V is given by

$$F(v) = \Pr[V \leq v] = \Pr[10,000e^R \leq v] = \Pr[R \leq \ln(v) - \ln(10,000)]$$

$$= \frac{1}{0.04} \int_{0.04}^{\ln(v) - \ln(10,000)} dr = \frac{1}{0.04} r \Big|_{0.04}^{\ln(v) - \ln(10,000)} = 25 \ln(v) - 25 \ln(10,000) - 1$$

$$= 25 \left[\ln\left(\frac{v}{10,000}\right) - 0.04 \right]$$

73. Solution: E

$$F(y) = \Pr[Y \leq y] = \Pr[10X^{0.8} \leq y] = \Pr\left[X \leq \left(\frac{Y}{10}\right)^{10/8}\right] = 1 - e^{-(Y/10)^{10/8}}$$

$$\text{Therefore, } f(y) = F'(y) = \frac{1}{8} \left(\frac{Y}{10}\right)^{1/4} e^{-(Y/10)^{5/4}}$$

74. Solution: E

First note $R = 10/T$. Then

$$F_R(r) = P[R \leq r] = P\left[\frac{10}{T} \leq r\right] = P\left[T \geq \frac{10}{r}\right] = 1 - F_T\left(\frac{10}{r}\right). \text{ Differentiating with respect to}$$

$$r \quad f_R(r) = F'_R(r) = d/dr \left(1 - F_T\left(\frac{10}{r}\right)\right) = -\left(\frac{d}{dt} F_T(t)\right)\left(\frac{-10}{r^2}\right)$$

$$\frac{d}{dt} F_T(t) = f_T(t) = \frac{1}{4} \text{ since } T \text{ is uniformly distributed on } [8, 12].$$

$$\text{Therefore } f_R(r) = \frac{-1}{4} \left(\frac{-10}{r^2}\right) = \frac{5}{2r^2}.$$

75. Solution: A

Let X and Y be the monthly profits of Company I and Company II, respectively. We are given that the pdf of X is f . Let us also take g to be the pdf of Y and take F and G to be the distribution functions corresponding to f and g . Then $G(y) = \Pr[Y \leq y] = \Pr[2X \leq y] = \Pr[X \leq y/2] = F(y/2)$ and $g(y) = G'(y) = d/dy F(y/2) = \frac{1}{2} F'(y/2) = \frac{1}{2} f(y/2)$.

76. Solution: A

First, observe that the distribution function of X is given by

$$F(x) = \int_1^x \frac{3}{t^4} dt = -\frac{1}{t^3} \Big|_1^x = 1 - \frac{1}{x^3}, \quad x > 1$$

Next, let X_1, X_2 , and X_3 denote the three claims made that have this distribution. Then if Y denotes the largest of these three claims, it follows that the distribution function of Y is given by

$$\begin{aligned} G(y) &= \Pr[X_1 \leq y] \Pr[X_2 \leq y] \Pr[X_3 \leq y] \\ &= \left(1 - \frac{1}{y^3}\right)^3, \quad y > 1 \end{aligned}$$

while the density function of Y is given by

$$g(y) = G'(y) = 3 \left(1 - \frac{1}{y^3}\right)^2 \left(\frac{3}{y^4}\right) = \left(\frac{9}{y^4}\right) \left(1 - \frac{1}{y^3}\right)^2, \quad y > 1$$

Therefore,

$$\begin{aligned}
 E[Y] &= \int_1^{\infty} \frac{9}{y^3} \left(1 - \frac{1}{y^3}\right)^2 dy = \int_1^{\infty} \frac{9}{y^3} \left(1 - \frac{2}{y^3} + \frac{1}{y^6}\right) dy \\
 &= \int_1^{\infty} \left(\frac{9}{y^3} - \frac{18}{y^6} + \frac{9}{y^9}\right) dy = \left[-\frac{9}{2y^2} + \frac{18}{5y^5} - \frac{9}{8y^8}\right]_1^{\infty} \\
 &= 9 \left[\frac{1}{2} - \frac{2}{5} + \frac{1}{8}\right] = 2.025 \text{ (in thousands)}
 \end{aligned}$$

77. Solution: D

$$\text{Prob.} = 1 - \int_1^2 \int_1^2 \frac{1}{8}(x+y) dx dy = 0.625$$

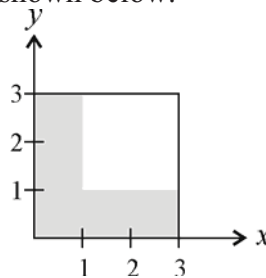
Note

$$\Pr[(X \leq 1) \cup (Y \leq 1)] = \Pr\left\{[(X > 1) \cap (Y > 1)]^c\right\} \quad (\text{De Morgan's Law})$$

$$\begin{aligned}
 &= 1 - \Pr[(X > 1) \cap (Y > 1)] = 1 - \int_1^2 \int_1^2 \frac{1}{8}(x+y) dx dy = 1 - \frac{1}{8} \int_1^2 \frac{1}{2}(x+y)^2 \Big|_1^2 dy \\
 &= 1 - \frac{1}{16} \int_1^2 [(y+2)^2 - (y+1)^2] dy = 1 - \frac{1}{48} [(y+2)^3 - (y+1)^3] \Big|_1^2 = 1 - \frac{1}{48} (64 - 27 - 27 + 8) \\
 &= 1 - \frac{18}{48} = \frac{30}{48} = 0.625
 \end{aligned}$$

78. Solution: B

That the device fails within the first hour means the joint density function must be integrated over the shaded region shown below.

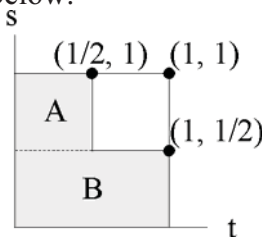


This evaluation is more easily performed by integrating over the unshaded region and subtracting from 1.

$$\begin{aligned}
& \Pr[(X < 1) \cup (Y < 1)] \\
&= 1 - \int_1^3 \int_1^3 \frac{x+y}{27} dx dy = 1 - \int_1^3 \left. \frac{x^2 + 2xy}{54} \right|_1^3 dy = 1 - \frac{1}{54} \int_1^3 (9 + 6y - 1 - 2y) dy \\
&= 1 - \frac{1}{54} \int_1^3 (8 + 4y) dy = 1 - \frac{1}{54} (8y + 2y^2) \Big|_1^3 = 1 - \frac{1}{54} (24 + 18 - 8 - 2) = 1 - \frac{32}{54} = \frac{11}{27} = 0.41
\end{aligned}$$

79. Solution: E

The domain of s and t is pictured below.



Note that the shaded region is the portion of the domain of s and t over which the device fails sometime during the first half hour. Therefore,

$$\Pr\left[\left(S \leq \frac{1}{2}\right) \cup \left(T \leq \frac{1}{2}\right)\right] = \int_0^{1/2} \int_{1/2}^1 f(s, t) ds dt + \int_0^1 \int_0^{1/2} f(s, t) ds dt$$

(where the first integral covers A and the second integral covers B).

80. Solution: C

By the central limit theorem, the total contributions are approximately normally distributed with mean $n\mu = (2025)(3125) = 6,328,125$ and standard deviation

$\sigma\sqrt{n} = 250\sqrt{2025} = 11,250$. From the tables, the 90th percentile for a standard normal random variable is 1.282. Letting p be the 90th percentile for total contributions,

$$\frac{p - n\mu}{\sigma\sqrt{n}} = 1.282, \text{ and so } p = n\mu + 1.282 \sigma\sqrt{n} = 6,328,125 + (1.282)(11,250) = 6,342,548.$$

81. Solution: C

Let X_1, \dots, X_{25} denote the 25 collision claims, and let $\bar{X} = \frac{1}{25}(X_1 + \dots + X_{25})$. We are given that each X_i ($i = 1, \dots, 25$) follows a normal distribution with mean 19,400 and standard deviation 5000. As a result \bar{X} also follows a normal distribution with mean 19,400 and standard deviation $\frac{1}{\sqrt{25}}(5000) = 1000$. We conclude that $P[\bar{X} > 20,000]$

$$= P\left[\frac{\bar{X} - 19,400}{1000} > \frac{20,000 - 19,400}{1000}\right] = P\left[\frac{\bar{X} - 19,400}{1000} > 0.6\right] = 1 - \Phi(0.6) = 1 - 0.7257 = 0.2743.$$

82. Solution: B

Let X_1, \dots, X_{1250} be the number of claims filed by each of the 1250 policyholders. We are given that each X_i follows a Poisson distribution with mean 2. It follows that $E[X_i] = \text{Var}[X_i] = 2$. Now we are interested in the random variable $S = X_1 + \dots + X_{1250}$. Assuming that the random variables are independent, we may conclude that S has an approximate normal distribution with $E[S] = \text{Var}[S] = (2)(1250) = 2500$.

Therefore $P[2450 < S < 2600] =$

$$P\left[\frac{2450 - 2500}{\sqrt{2500}} < \frac{S - 2500}{\sqrt{2500}} < \frac{2600 - 2500}{\sqrt{2500}}\right] = P\left[-1 < \frac{S - 2500}{50} < 2\right]$$
$$= P\left[\frac{S - 2500}{50} < 2\right] - P\left[\frac{S - 2500}{50} < -1\right]$$

Then using the normal approximation with $Z = \frac{S - 2500}{50}$, we have $P[2450 < S < 2600] \approx P[Z < 2] - P[Z > 1] = P[Z < 2] + P[Z < 1] - 1 \approx 0.9773 + 0.8413 - 1 = 0.8186$.

83. Solution: B

Let X_1, \dots, X_n denote the life spans of the n light bulbs purchased. Since these random variables are independent and normally distributed with mean 3 and variance 1, the random variable $S = X_1 + \dots + X_n$ is also normally distributed with mean

$$\mu = 3n$$

and standard deviation

$$\sigma = \sqrt{n}$$

Now we want to choose the smallest value for n such that

$$0.9772 \leq \Pr[S > 40] = \Pr\left[\frac{S - 3n}{\sqrt{n}} > \frac{40 - 3n}{\sqrt{n}}\right]$$

This implies that n should satisfy the following inequality:

$$-2 \geq \frac{40-3n}{\sqrt{n}}$$

To find such an n , let's solve the corresponding equation for n :

$$-2 = \frac{40-3n}{\sqrt{n}}$$

$$-2\sqrt{n} = 40 - 3n$$

$$3n - 2\sqrt{n} - 40 = 0$$

$$(3\sqrt{n} + 10)(\sqrt{n} - 4) = 0$$

$$\sqrt{n} = 4$$

$$n = 16$$

84. Solution: B

Observe that

$$E[X + Y] = E[X] + E[Y] = 50 + 20 = 70$$

$$Var[X + Y] = Var[X] + Var[Y] + 2 Cov[X, Y] = 50 + 30 + 20 = 100$$

for a randomly selected person. It then follows from the Central Limit Theorem that T is approximately normal with mean

$$E[T] = 100(70) = 7000$$

and variance

$$Var[T] = 100(100) = 100^2$$

Therefore,

$$\begin{aligned} \Pr[T < 7100] &= \Pr\left[\frac{T - 7000}{100} < \frac{7100 - 7000}{100}\right] \\ &= \Pr[Z < 1] = 0.8413 \end{aligned}$$

where Z is a standard normal random variable.

85. Solution: B

Denote the policy premium by P . Since x is exponential with parameter 1000, it follows from what we are given that $E[X] = 1000$, $\text{Var}[X] = 1,000,000$, $\sqrt{\text{Var}[X]} = 1000$ and $P = 100 + E[X] = 1,100$. Now if 100 policies are sold, then Total Premium Collected = $100(1,100) = 110,000$

Moreover, if we denote total claims by S , and assume the claims of each policy are independent of the others then $E[S] = 100 E[X] = (100)(1000)$ and $\text{Var}[S] = 100 \text{Var}[X] = (100)(1,000,000)$. It follows from the Central Limit Theorem that S is approximately normally distributed with mean 100,000 and standard deviation = 10,000. Therefore,

$$P[S \geq 110,000] = 1 - P[S \leq 110,000] = 1 - P\left[Z \leq \frac{110,000 - 100,000}{10,000}\right] = 1 - P[Z \leq 1] = 1 - 0.841 \approx 0.159.$$

86. Solution: E

Let X_1, \dots, X_{100} denote the number of pensions that will be provided to each new recruit. Now under the assumptions given,

$$X_i = \begin{cases} 0 & \text{with probability } 1 - 0.4 = 0.6 \\ 1 & \text{with probability } (0.4)(0.25) = 0.1 \\ 2 & \text{with probability } (0.4)(0.75) = 0.3 \end{cases}$$

for $i = 1, \dots, 100$. Therefore,

$$E[X_i] = (0)(0.6) + (1)(0.1) + (2)(0.3) = 0.7,$$

$$E[X_i^2] = (0)^2(0.6) + (1)^2(0.1) + (2)^2(0.3) = 1.3, \text{ and}$$

$$\text{Var}[X_i] = E[X_i^2] - \{E[X_i]\}^2 = 1.3 - (0.7)^2 = 0.81$$

Since X_1, \dots, X_{100} are assumed by the consulting actuary to be independent, the Central Limit Theorem then implies that $S = X_1 + \dots + X_{100}$ is approximately normally distributed with mean

$$E[S] = E[X_1] + \dots + E[X_{100}] = 100(0.7) = 70$$

and variance

$$\text{Var}[S] = \text{Var}[X_1] + \dots + \text{Var}[X_{100}] = 100(0.81) = 81$$

Consequently,

$$\begin{aligned} \Pr[S \leq 90.5] &= \Pr\left[\frac{S - 70}{9} \leq \frac{90.5 - 70}{9}\right] \\ &= \Pr[Z \leq 2.28] \\ &= 0.99 \end{aligned}$$

87. Solution: D

Let X denote the difference between true and reported age. We are given X is uniformly distributed on $(-2.5, 2.5)$. That is, X has pdf $f(x) = 1/5$, $-2.5 < x < 2.5$. It follows that $\mu_x = E[X] = 0$

$$\sigma_x^2 = \text{Var}[X] = E[X^2] = \int_{-2.5}^{2.5} \frac{x^2}{5} dx = \frac{x^3}{15} \Big|_{-2.5}^{2.5} = \frac{2(2.5)^3}{15} = 2.083$$

$$\sigma_x = 1.443$$

Now \bar{X}_{48} , the difference between the means of the true and rounded ages, has a

distribution that is approximately normal with mean 0 and standard deviation $\frac{1.443}{\sqrt{48}} =$

0.2083. Therefore,

$$\begin{aligned} P\left[-\frac{1}{4} \leq \bar{X}_{48} \leq \frac{1}{4}\right] &= P\left[\frac{-0.25}{0.2083} \leq Z \leq \frac{0.25}{0.2083}\right] = P[-1.2 \leq Z \leq 1.2] = P[Z \leq 1.2] - P[Z \leq -1.2] \\ &= P[Z \leq 1.2] - 1 + P[Z \leq 1.2] = 2P[Z \leq 1.2] - 1 = 2(0.8849) - 1 = 0.77. \end{aligned}$$

88. Solution: C

Let X denote the waiting time for a first claim from a good driver, and let Y denote the waiting time for a first claim from a bad driver. The problem statement implies that the respective distribution functions for X and Y are

$$F(x) = 1 - e^{-x/6}, \quad x > 0 \quad \text{and}$$

$$G(y) = 1 - e^{-y/3}, \quad y > 0$$

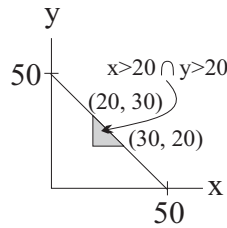
Therefore,

$$\begin{aligned} \Pr[(X \leq 3) \cap (Y \leq 2)] &= \Pr[X \leq 3] \Pr[Y \leq 2] \\ &= F(3)G(2) = (1 - e^{-1/2})(1 - e^{-2/3}) = 1 - e^{-2/3} - e^{-1/2} + e^{-7/6} \end{aligned}$$

89. Solution: B

$$\text{We are given that } f(x, y) = \begin{cases} \frac{6}{125,000}(50 - x - y) & \text{for } 0 < x < 50 - y < 50 \\ 0 & \text{otherwise} \end{cases}$$

and we want to determine $P[X > 20 \cap Y > 20]$. In order to determine integration limits, consider the following diagram:



$$\text{We conclude that } P[X > 20 \cap Y > 20] = \frac{6}{125,000} \int_{20}^{30} \int_{20}^{50-x} (50 - x - y) \, dy \, dx.$$

90. Solution: C

Let T_1 be the time until the next Basic Policy claim, and let T_2 be the time until the next Deluxe policy claim. Then the joint pdf of T_1 and T_2 is

$$f(t_1, t_2) = \left(\frac{1}{2} e^{-t_1/2} \right) \left(\frac{1}{3} e^{-t_2/3} \right) = \frac{1}{6} e^{-t_1/2} e^{-t_2/3}, \quad 0 < t_1 < \infty, \quad 0 < t_2 < \infty \text{ and we need to find}$$

$$\begin{aligned} P[T_2 < T_1] &= \int_0^{\infty} \int_0^{t_1} \frac{1}{6} e^{-t_1/2} e^{-t_2/3} \, dt_2 \, dt_1 = \int_0^{\infty} \left[-\frac{1}{2} e^{-t_1/2} e^{-t_2/3} \right]_0^{t_1} dt_1 \\ &= \int_0^{\infty} \left[\frac{1}{2} e^{-t_1/2} - \frac{1}{2} e^{-t_1/2} e^{-t_1/3} \right] dt_1 = \int_0^{\infty} \left[\frac{1}{2} e^{-t_1/2} - \frac{1}{2} e^{-5t_1/6} \right] dt_1 = \left[-e^{-t_1/2} + \frac{3}{5} e^{-5t_1/6} \right]_0^{\infty} = 1 - \frac{3}{5} = \frac{2}{5} \\ &= 0.4. \end{aligned}$$

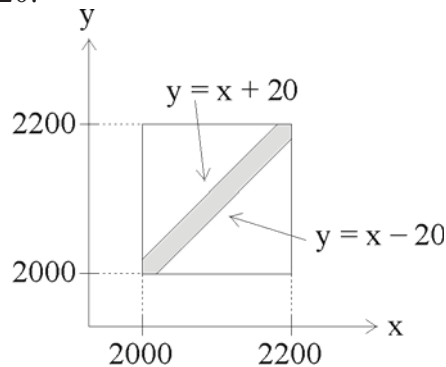
91. Solution: D

We want to find $P[X + Y > 1]$. To this end, note that $P[X + Y > 1]$

$$\begin{aligned} &= \int_0^1 \int_{1-x}^2 \left[\frac{2x+2-y}{4} \right] dy \, dx = \int_0^1 \left[\frac{1}{2} xy + \frac{1}{2} y - \frac{1}{8} y^2 \right]_{1-x}^2 dx \\ &= \int_0^1 \left[x + 1 - \frac{1}{2} - \frac{1}{2} x(1-x) - \frac{1}{2} (1-x) + \frac{1}{8} (1-x)^2 \right] dx = \int_0^1 \left[x + \frac{1}{2} x^2 + \frac{1}{8} - \frac{1}{4} x + \frac{1}{8} x^2 \right] dx \\ &= \int_0^1 \left[\frac{5}{8} x^2 + \frac{3}{4} x + \frac{1}{8} \right] dx = \left[\frac{5}{24} x^3 + \frac{3}{8} x^2 + \frac{1}{8} x \right]_0^1 = \frac{5}{24} + \frac{3}{8} + \frac{1}{8} = \frac{17}{24} \end{aligned}$$

92. Solution: B

Let X and Y denote the two bids. Then the graph below illustrates the region over which X and Y differ by less than 20:



Based on the graph and the uniform distribution:

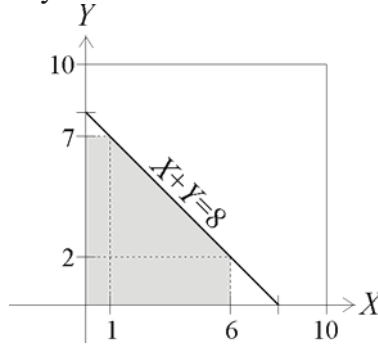
$$\begin{aligned}\Pr[|X - Y| < 20] &= \frac{\text{Shaded Region Area}}{(2200 - 2000)^2} = \frac{200^2 - 2 \cdot \frac{1}{2}(180)^2}{200^2} \\ &= 1 - \frac{180^2}{200^2} = 1 - (0.9)^2 = 0.19\end{aligned}$$

More formally (still using symmetry)

$$\begin{aligned}\Pr[|X - Y| < 20] &= 1 - \Pr[|X - Y| \geq 20] = 1 - 2 \Pr[X - Y \geq 20] \\ &= 1 - 2 \int_{2020}^{2200} \int_{2000}^{x-20} \frac{1}{200^2} dy dx = 1 - 2 \int_{2020}^{2200} \frac{1}{200^2} y \Big|_{2000}^{x-20} dx \\ &= 1 - \frac{2}{200^2} \int_{2020}^{2200} (x - 20 - 2000) dx = 1 - \frac{1}{200^2} (x - 2020)^2 \Big|_{2020}^{2200} \\ &= 1 - \left(\frac{180}{200}\right)^2 = 0.19\end{aligned}$$

93. Solution: C

Define X and Y to be loss amounts covered by the policies having deductibles of 1 and 2, respectively. The shaded portion of the graph below shows the region over which the total benefit paid to the family does not exceed 5:



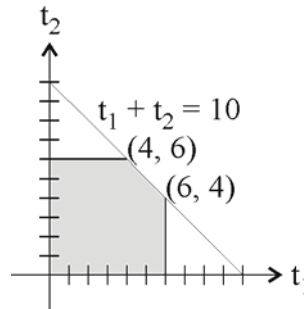
We can also infer from the graph that the uniform random variables X and Y have joint density function $f(x, y) = \frac{1}{100}$, $0 < x < 10$, $0 < y < 10$

We could integrate f over the shaded region in order to determine the desired probability. However, since X and Y are uniform random variables, it is simpler to determine the portion of the 10×10 square that is shaded in the graph above. That is,

$$\begin{aligned} & \Pr(\text{Total Benefit Paid Does not Exceed } 5) \\ &= \Pr(0 < X < 6, 0 < Y < 2) + \Pr(0 < X < 1, 2 < Y < 7) + \Pr(1 < X < 6, 2 < Y < 8 - X) \\ &= \frac{(6)(2)}{100} + \frac{(1)(5)}{100} + \frac{(1/2)(5)(5)}{100} = \frac{12}{100} + \frac{5}{100} + \frac{12.5}{100} = 0.295 \end{aligned}$$

94. Solution: C

Let $f(t_1, t_2)$ denote the joint density function of T_1 and T_2 . The domain of f is pictured below:



Now the area of this domain is given by

$$A = 6^2 - \frac{1}{2}(6-4)^2 = 36 - 2 = 34$$

$$\text{Consequently, } f(t_1, t_2) = \begin{cases} \frac{1}{34} & , \quad 0 < t_1 < 6, \quad 0 < t_2 < 6, \quad t_1 + t_2 < 10 \\ 0 & \text{elsewhere} \end{cases}$$

and

$$\begin{aligned} E[T_1 + T_2] &= E[T_1] + E[T_2] = 2E[T_1] \quad (\text{due to symmetry}) \\ &= 2 \left\{ \int_0^4 t_1 \int_0^6 \frac{1}{34} dt_2 dt_1 + \int_4^6 t_1 \int_0^{10-t_1} \frac{1}{34} dt_2 dt_1 \right\} = 2 \left\{ \int_0^4 t_1 \left[\frac{t_2}{34} \Big|_0^6 \right] dt_1 + \int_4^6 t_1 \left[\frac{t_2}{34} \Big|_0^{10-t_1} \right] dt_1 \right\} \\ &= 2 \left\{ \int_0^4 \frac{3t_1}{17} dt_1 + \int_4^6 \frac{1}{34} (10t_1 - t_1^2) dt_1 \right\} = 2 \left\{ \frac{3t_1^2}{34} \Big|_0^4 + \frac{1}{34} \left(5t_1^2 - \frac{1}{3} t_1^3 \right) \Big|_4^6 \right\} \\ &= 2 \left\{ \frac{24}{17} + \frac{1}{34} \left[180 - 72 - 80 + \frac{64}{3} \right] \right\} = 5.7 \end{aligned}$$

95. Solution: E

$$\begin{aligned} M(t_1, t_2) &= E[e^{t_1 W + t_2 Z}] = E[e^{t_1(X+Y) + t_2(Y-X)}] = E[e^{(t_1-t_2)X} e^{(t_1+t_2)Y}] \\ &= E[e^{(t_1-t_2)X}] E[e^{(t_1+t_2)Y}] = e^{\frac{1}{2}(t_1-t_2)^2} e^{\frac{1}{2}(t_1+t_2)^2} = e^{\frac{1}{2}(t_1^2 - 2t_1t_2 + t_2^2)} e^{\frac{1}{2}(t_1^2 + 2t_1t_2 + t_2^2)} = e^{t_1^2 + t_2^2} \end{aligned}$$

96. Solution: E

Observe that the bus driver collect $21 \times 50 = 1050$ for the 21 tickets he sells. However, he may be required to refund 100 to one passenger if all 21 ticket holders show up. Since passengers show up or do not show up independently of one another, the probability that all 21 passengers will show up is $(1 - 0.02)^{21} = (0.98)^{21} = 0.65$. Therefore, the tour operator's expected revenue is $1050 - (100)(0.65) = 985$.

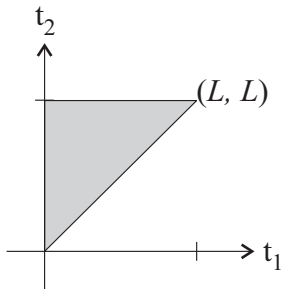
97. Solution: C

We are given $f(t_1, t_2) = 2/L^2, 0 \leq t_1 \leq t_2 \leq L$.

$$\text{Therefore, } E[T_1^2 + T_2^2] = \int_0^L \int_0^{t_2} (t_1^2 + t_2^2) \frac{2}{L^2} dt_1 dt_2 =$$

$$\frac{2}{L^2} \left\{ \int_0^L \left[\frac{t_1^3}{3} + t_2^2 t_1 \right]_0^{t_2} dt_2 \right\} = \frac{2}{L^2} \left\{ \int_0^L \left(\frac{t_2^3}{3} + t_2^3 \right) dt_2 \right\}$$

$$= \frac{2}{L^2} \int_0^L \frac{4}{3} t_2^3 dt_2 = \frac{2}{L^2} \left[\frac{t_2^4}{3} \right]_0^L = \frac{2}{3} L^2$$



98. Solution: A

Let $g(y)$ be the probability function for $Y = X_1 X_2 X_3$. Note that $Y = 1$ if and only if $X_1 = X_2 = X_3 = 1$. Otherwise, $Y = 0$. Since $P[Y = 1] = P[X_1 = 1 \cap X_2 = 1 \cap X_3 = 1] = P[X_1 = 1] P[X_2 = 1] P[X_3 = 1] = (2/3)^3 = 8/27$.

$$\text{We conclude that } g(y) = \begin{cases} \frac{19}{27} & \text{for } y = 0 \\ \frac{8}{27} & \text{for } y = 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\text{and } M(t) = E[e^{y_t}] = \frac{19}{27} + \frac{8}{27} e^t$$

99. Solution: C

We use the relationships $\text{Var}(aX + b) = a^2 \text{Var}(X)$, $\text{Cov}(aX, bY) = ab \text{Cov}(X, Y)$, and $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2 \text{Cov}(X, Y)$. First we observe

$17,000 = \text{Var}(X + Y) = 5000 + 10,000 + 2 \text{Cov}(X, Y)$, and so $\text{Cov}(X, Y) = 1000$.

We want to find $\text{Var}[(X + 100) + 1.1Y] = \text{Var}[(X + 1.1Y) + 100]$

$= \text{Var}[X + 1.1Y] = \text{Var} X + \text{Var}[(1.1)Y] + 2 \text{Cov}(X, 1.1Y)$

$= \text{Var} X + (1.1)^2 \text{Var} Y + 2(1.1) \text{Cov}(X, Y) = 5000 + 12,100 + 2200 = 19,300$.

100. Solution: B

Note

$$P(X = 0) = 1/6$$

$$P(X = 1) = 1/12 + 1/6 = 3/12$$

$$P(X = 2) = 1/12 + 1/3 + 1/6 = 7/12$$

$$E[X] = (0)(1/6) + (1)(3/12) + (2)(7/12) = 17/12$$

$$E[X^2] = (0)^2(1/6) + (1)^2(3/12) + (2)^2(7/12) = 31/12$$

$$\text{Var}[X] = 31/12 - (17/12)^2 = 0.58$$

101. Solution: D

Note that due to the independence of X and Y

$$\text{Var}(Z) = \text{Var}(3X - Y - 5) = \text{Var}(3X) + \text{Var}(Y) = 3^2 \text{Var}(X) + \text{Var}(Y) = 9(1) + 2 = 11$$

102. Solution: E

Let X and Y denote the times that the two backup generators can operate. Now the variance of an exponential random variable with mean β is β^2 . Therefore,

$$\text{Var}[X] = \text{Var}[Y] = 10^2 = 100$$

Then assuming that X and Y are independent, we see

$$\text{Var}[X+Y] = \text{Var}[X] + \text{Var}[Y] = 100 + 100 = 200$$

103. Solution: E

Let X_1, X_2 , and X_3 denote annual loss due to storm, fire, and theft, respectively. In addition, let $Y = \text{Max}(X_1, X_2, X_3)$.

Then

$$\begin{aligned}\Pr[Y > 3] &= 1 - \Pr[Y \leq 3] = 1 - \Pr[X_1 \leq 3] \Pr[X_2 \leq 3] \Pr[X_3 \leq 3] \\ &= 1 - (1 - e^{-3}) \left(1 - e^{-3/1.5}\right) \left(1 - e^{-3/2.4}\right) \quad * \\ &= 1 - (1 - e^{-3}) (1 - e^{-2}) \left(1 - e^{-5/4}\right) \\ &= 0.414\end{aligned}$$

* Uses that if X has an exponential distribution with mean μ

$$\Pr(X \leq x) = 1 - \Pr(X \geq x) = 1 - \int_x^\infty \frac{1}{\mu} e^{-t/\mu} dt = 1 - \left(-e^{-t/\mu}\right) \Big|_x^\infty = 1 - e^{-x/\mu}$$

104. Solution: B

Let us first determine k :

$$1 = \int_0^1 \int_0^1 kx dx dy = \int_0^1 \frac{1}{2} kx^2 \Big|_0^1 dy = \int_0^1 \frac{k}{2} dy = \frac{k}{2}$$

$$k = 2$$

Then

$$E[X] = \int_0^1 \int_0^1 2x^2 dy dx = \int_0^1 2x^2 dx = \frac{2}{3} x^3 \Big|_0^1 = \frac{2}{3}$$

$$E[Y] = \int_0^1 \int_0^1 2x dx dy = \int_0^1 y dy = \frac{1}{2} y^2 \Big|_0^1 = \frac{1}{2}$$

$$\begin{aligned}E[XY] &= \int_0^1 \int_0^1 2x^2 y dx dy = \int_0^1 \frac{2}{3} x^3 y \Big|_0^1 dy = \int_0^1 \frac{2}{3} y dy \\ &= \frac{2}{6} y^2 \Big|_0^1 = \frac{2}{6} = \frac{1}{3}\end{aligned}$$

$$\text{Cov}[X, Y] = E[XY] - E[X]E[Y] = \frac{1}{3} - \left(\frac{2}{3}\right)\left(\frac{1}{2}\right) = \frac{1}{3} - \frac{1}{3} = 0$$

(Alternative Solution)

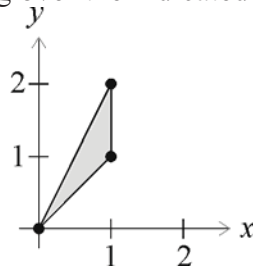
Define $g(x) = kx$ and $h(y) = 1$. Then

$$f(x, y) = g(x)h(y)$$

In other words, $f(x, y)$ can be written as the product of a function of x alone and a function of y alone. It follows that X and Y are independent. Therefore, $\text{Cov}[X, Y] = 0$.

105. Solution: A

The calculation requires integrating over the indicated region.



$$E(X) = \int_0^1 \int_x^{2x} \frac{8}{3} x^2 y \, dy \, dx = \int_0^1 \frac{4}{3} x^2 y^2 \Big|_x^{2x} \, dx = \int_0^1 \frac{4}{3} x^2 (4x^2 - x^2) \, dx = \int_0^1 4x^4 \, dx = \frac{4}{5} x^5 \Big|_0^1 = \frac{4}{5}$$

$$E(Y) = \int_0^1 \int_x^{2x} \frac{8}{3} x y^2 \, dy \, dx = \int_0^1 \frac{8}{9} x y^3 \Big|_x^{2x} \, dx = \int_0^1 \frac{8}{9} x (8x^3 - x^3) \, dx = \int_0^1 \frac{56}{9} x^4 \, dx = \frac{56}{45} x^5 \Big|_0^1 = \frac{56}{45}$$

$$E(XY) = \int_0^1 \int_x^{2x} \frac{8}{3} x^2 y^2 \, dy \, dx = \int_0^1 \frac{8}{9} x^2 y^3 \Big|_x^{2x} \, dx = \int_0^1 \frac{8}{9} x^2 (8x^3 - x^3) \, dx = \int_0^1 \frac{56}{9} x^5 \, dx = \frac{56}{54} x^6 \Big|_0^1 = \frac{28}{27}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{28}{27} - \left(\frac{56}{45}\right)\left(\frac{4}{5}\right) = 0.04$$

106. Solution: C

The joint pdf of X and Y is $f(x, y) = f_2(y|x) f_1(x)$

$= (1/x)(1/12)$, $0 < y < x$, $0 < x < 12$.

Therefore,

$$E[X] = \int_0^{12} \int_0^x x \cdot \frac{1}{12x} \, dy \, dx = \int_0^{12} \frac{y}{12} \Big|_0^x \, dx = \int_0^{12} \frac{x}{12} \, dx = \frac{x^2}{24} \Big|_0^{12} = 6$$

$$E[Y] = \int_0^{12} \int_0^x \frac{y}{12x} \, dy \, dx = \int_0^{12} \left[\frac{y^2}{24x} \right]_0^x \, dx = \int_0^{12} \frac{x}{24} \, dx = \frac{x^2}{48} \Big|_0^{12} = \frac{144}{48} = 3$$

$$E[XY] = \int_0^{12} \int_0^x \frac{y}{12} \, dy \, dx = \int_0^{12} \left[\frac{y^2}{24} \right]_0^x \, dx = \int_0^{12} \frac{x^2}{24} \, dx = \frac{x^3}{72} \Big|_0^{12} = \frac{(12)^3}{72} = 24$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 24 - (3)(6) = 24 - 18 = 6.$$

107. Solution: A

$$\begin{aligned}
 \text{Cov}(C_1, C_2) &= \text{Cov}(X + Y, X + 1.2Y) \\
 &= \text{Cov}(X, X) + \text{Cov}(Y, X) + \text{Cov}(X, 1.2Y) + \text{Cov}(Y, 1.2Y) \\
 &= \text{Var } X + \text{Cov}(X, Y) + 1.2\text{Cov}(X, Y) + 1.2\text{Var } Y \\
 &= \text{Var } X + 2.2\text{Cov}(X, Y) + 1.2\text{Var } Y
 \end{aligned}$$

$$\text{Var } X = E(X^2) - (E(X))^2 = 27.4 - 5^2 = 2.4$$

$$\text{Var } Y = E(Y^2) - (E(Y))^2 = 51.4 - 7^2 = 2.4$$

$$\text{Var}(X + Y) = \text{Var } X + \text{Var } Y + 2\text{Cov}(X, Y)$$

$$\text{Cov}(X, Y) = \frac{1}{2}(\text{Var}(X + Y) - \text{Var } X - \text{Var } Y) = \frac{1}{2}(8 - 2.4 - 2.4) = 1.6$$

$$\text{Cov}(C_1, C_2) = 2.4 + 2.2(1.6) + 1.2(2.4) = 8.8$$

107. Alternate solution:

We are given the following information:

$$C_1 = X + Y$$

$$C_2 = X + 1.2Y$$

$$E[X] = 5$$

$$E[X^2] = 27.4$$

$$E[Y] = 7$$

$$E[Y^2] = 51.4$$

$$\text{Var}[X + Y] = 8$$

Now we want to calculate

$$\begin{aligned}
 \text{Cov}(C_1, C_2) &= \text{Cov}(X + Y, X + 1.2Y) \\
 &= E[(X + Y)(X + 1.2Y)] - E[X + Y] \cdot E[X + 1.2Y] \\
 &= E[X^2 + 2.2XY + 1.2Y^2] - (E[X] + E[Y])(E[X] + 1.2E[Y]) \\
 &= E[X^2] + 2.2E[XY] + 1.2E[Y^2] - (5 + 7)(5 + (1.2)7) \\
 &= 27.4 + 2.2E[XY] + 1.2(51.4) - (12)(13.4) \\
 &= 2.2E[XY] - 71.72
 \end{aligned}$$

Therefore, we need to calculate $E[XY]$ first. To this end, observe

$$\begin{aligned}
8 = \text{Var}[X + Y] &= E[(X + Y)^2] - (E[X + Y])^2 \\
&= E[X^2 + 2XY + Y^2] - (E[X] + E[Y])^2 \\
&= E[X^2] + 2E[XY] + E[Y^2] - (5 + 7)^2 \\
&= 27.4 + 2E[XY] + 51.4 - 144 \\
&= 2E[XY] - 65.2
\end{aligned}$$

$$E[XY] = (8 + 65.2)/2 = 36.6$$

$$\text{Finally, } \text{Cov}(C_1, C_2) = 2.2(36.6) - 71.72 = 8.8$$

108. Solution: A

The joint density of T_1 and T_2 is given by

$$f(t_1, t_2) = e^{-t_1} e^{-t_2}, \quad t_1 > 0, \quad t_2 > 0$$

Therefore,

$$\begin{aligned}
\Pr[X \leq x] &= \Pr[2T_1 + T_2 \leq x] \\
&= \int_0^x \int_0^{\frac{1}{2}(x-t_2)} e^{-t_1} e^{-t_2} dt_1 dt_2 = \int_0^x e^{-t_2} \left[-e^{-t_1} \right]_0^{\frac{1}{2}(x-t_2)} dt_2 \\
&= \int_0^x e^{-t_2} \left[1 - e^{-\frac{1}{2}x + \frac{1}{2}t_2} \right] dt_2 = \int_0^x \left(e^{-t_2} - e^{-\frac{1}{2}x} e^{-\frac{1}{2}t_2} \right) dt_2 \\
&= \left[-e^{-t_2} + 2e^{-\frac{1}{2}x} e^{-\frac{1}{2}t_2} \right]_0^x = -e^{-x} + 2e^{-\frac{1}{2}x} e^{-\frac{1}{2}x} + 1 - 2e^{-\frac{1}{2}x} \\
&= 1 - e^{-x} + 2e^{-x} - 2e^{-\frac{1}{2}x} = 1 - 2e^{-\frac{1}{2}x} + e^{-x}, \quad x > 0
\end{aligned}$$

It follows that the density of X is given by

$$g(x) = \frac{d}{dx} \left[1 - 2e^{-\frac{1}{2}x} + e^{-x} \right] = e^{-\frac{1}{2}x} - e^{-x}, \quad x > 0$$

109. Solution: B

Let

u be annual claims,

v be annual premiums,

$g(u, v)$ be the joint density function of U and V ,

$f(x)$ be the density function of X , and

$F(x)$ be the distribution function of X .

Then since U and V are independent,

$$g(u, v) = (e^{-u}) \left(\frac{1}{2} e^{-v/2} \right) = \frac{1}{2} e^{-u} e^{-v/2}, \quad 0 < u < \infty, \quad 0 < v < \infty$$

and

$$\begin{aligned} F(x) &= \Pr[X \leq x] = \Pr\left[\frac{u}{v} \leq x\right] = \Pr[U \leq Vx] \\ &= \int_0^\infty \int_0^{vx} g(u, v) du dv = \int_0^\infty \int_0^{vx} \frac{1}{2} e^{-u} e^{-v/2} du dv \\ &= \int_0^\infty \left[-\frac{1}{2} e^{-u} e^{-v/2} \right]_0^{vx} dv = \int_0^\infty \left(-\frac{1}{2} e^{-vx} e^{-v/2} + \frac{1}{2} e^{-v/2} \right) dv \\ &= \int_0^\infty \left(-\frac{1}{2} e^{-v(x+1/2)} + \frac{1}{2} e^{-v/2} \right) dv \\ &= \left[\frac{1}{2x+1} e^{-v(x+1/2)} - e^{-v/2} \right]_0^\infty = -\frac{1}{2x+1} + 1 \end{aligned}$$

$$\text{Finally, } f(x) = F'(x) = \frac{2}{(2x+1)^2}$$

110. Solution: C

Note that the conditional density function

$$f\left(y \middle| x = \frac{1}{3}\right) = \frac{f(1/3, y)}{f_x(1/3)}, \quad 0 < y < \frac{2}{3},$$

$$f_x\left(\frac{1}{3}\right) = \int_0^{2/3} 24(1/3)y dy = \int_0^{2/3} 8y dy = 4y^2 \Big|_0^{2/3} = \frac{16}{9}$$

$$\text{It follows that } f\left(y \middle| x = \frac{1}{3}\right) = \frac{9}{16} f(1/3, y) = \frac{9}{2} y, \quad 0 < y < \frac{2}{3}$$

$$\text{Consequently, } \Pr[Y < X | X = 1/3] = \int_0^{1/3} \frac{9}{2} y dy = \frac{9}{4} y^2 \Big|_0^{1/3} = \frac{1}{4}$$

111. Solution: E

$$\Pr[1 < Y < 3 | X = 2] = \int_1^3 \frac{f(2, y)}{f_x(2)} dy$$

$$f(2, y) = \frac{2}{4(2-1)} y^{-(4-1)/2-1} = \frac{1}{2} y^{-3}$$

$$f_x(2) = \int_1^\infty \frac{1}{2} y^{-3} dy = -\frac{1}{4} y^{-2} \Big|_1^\infty = \frac{1}{4}$$

$$\text{Finally, } \Pr[1 < Y < 3 | X = 2] = \frac{\int_1^3 \frac{1}{2} y^{-3} dy}{\frac{1}{4}} = -y^{-2} \Big|_1^3 = 1 - \frac{1}{9} = \frac{8}{9}$$

112. Solution: D

We are given that the joint pdf of X and Y is $f(x, y) = 2(x+y)$, $0 < y < x < 1$.

$$\text{Now } f_x(x) = \int_0^x (2x + 2y) dy = \left[2xy + y^2 \right]_0^x = 2x^2 + x^2 = 3x^2, 0 < x < 1$$

$$\text{so } f(y|x) = \frac{f(x, y)}{f_x(x)} = \frac{2(x+y)}{3x^2} = \frac{2}{3} \left(\frac{1}{x} + \frac{y}{x^2} \right), 0 < y < x$$

$$f(y|x = 0.10) = \frac{2}{3} \left[\frac{1}{0.1} + \frac{y}{0.01} \right] = \frac{2}{3} [10 + 100y], 0 < y < 0.10$$

$$P[Y < 0.05 | X = 0.10] = \int_0^{0.05} \frac{2}{3} [10 + 100y] dy = \left[\frac{20}{3} y + \frac{100}{3} y^2 \right]_0^{0.05} = \frac{1}{3} + \frac{1}{12} = \frac{5}{12} = 0.4167.$$

113. Solution: E

Let

W = event that wife survives at least 10 years

H = event that husband survives at least 10 years

B = benefit paid

P = profit from selling policies

$$\text{Then } \Pr[H] = P[H \cap W] + \Pr[H \cap W^c] = 0.96 + 0.01 = 0.97$$

and

$$\Pr[W | H] = \frac{\Pr[W \cap H]}{\Pr[H]} = \frac{0.96}{0.97} = 0.9897$$

$$\Pr[W^c | H] = \frac{\Pr[H \cap W^c]}{\Pr[H]} = \frac{0.01}{0.97} = 0.0103$$

It follows that

$$\begin{aligned} E[P] &= E[1000 - B] = 1000 - E[B] = 1000 - \{(0)\Pr[W | H] + (10,000)\Pr[W^c | H]\} \\ &= 1000 - 10,000(0.0103) = 1000 - 103 = 897 \end{aligned}$$

114. Solution: C

Note that

$$\begin{aligned} P(Y = 0 | X = 1) &= \frac{P(X = 1, Y = 0)}{P(X = 1)} = \frac{P(X = 1, Y = 0)}{P(X = 1, Y = 0) + P(X = 1, Y = 1)} = \frac{0.05}{0.05 + 0.125} \\ &= 0.286 \end{aligned}$$

$$P(Y = 1 | X = 1) = 1 - P(Y = 0 | X = 1) = 1 - 0.286 = 0.714$$

$$\text{Therefore, } E(Y | X = 1) = (0)P(Y = 0 | X = 1) + (1)P(Y = 1 | X = 1) = (1)(0.714) = 0.714$$

$$E(Y^2 | X = 1) = (0)^2 P(Y = 0 | X = 1) + (1)^2 P(Y = 1 | X = 1) = 0.714$$

$$\text{Var}(Y | X = 1) = E(Y^2 | X = 1) - [E(Y | X = 1)]^2 = 0.714 - (0.714)^2 = 0.20$$

115. Solution: A

Let $f_1(x)$ denote the marginal density function of X . Then

$$f_1(x) = \int_x^{x+1} 2xy dy = 2xy|_x^{x+1} = 2x(x+1-x) = 2x, \quad 0 < x < 1$$

Consequently,

$$f(y|x) = \frac{f(x,y)}{f_1(x)} = \begin{cases} 1 & \text{if: } x < y < x+1 \\ 0 & \text{otherwise} \end{cases}$$

$$E[Y|X] = \int_x^{x+1} y dy = \frac{1}{2} y^2|_x^{x+1} = \frac{1}{2} (x+1)^2 - \frac{1}{2} x^2 = \frac{1}{2} x^2 + x + \frac{1}{2} - \frac{1}{2} x^2 = x + \frac{1}{2}$$

$$E[Y^2|X] = \int_x^{x+1} y^2 dy = \frac{1}{3} y^3|_x^{x+1} = \frac{1}{3} (x+1)^3 - \frac{1}{3} x^3$$

$$= \frac{1}{3} x^3 + x^2 + x + \frac{1}{3} - \frac{1}{3} x^3 = x^2 + x + \frac{1}{3}$$

$$\text{Var}[Y|X] = E[Y^2|X] - \{E[Y|X]\}^2 = x^2 + x + \frac{1}{3} - \left(x + \frac{1}{2}\right)^2$$

$$= x^2 + x + \frac{1}{3} - x^2 - x - \frac{1}{4} = \frac{1}{12}$$

116. Solution: D

Denote the number of tornadoes in counties P and Q by N_P and N_Q , respectively. Then

$$E[N_Q|N_P = 0]$$

$$= [(0)(0.12) + (1)(0.06) + (2)(0.05) + 3(0.02)] / [0.12 + 0.06 + 0.05 + 0.02] = 0.88$$

$$E[N_Q^2|N_P = 0]$$

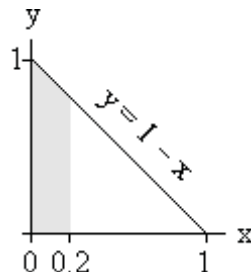
$$= [(0)^2(0.12) + (1)^2(0.06) + (2)^2(0.05) + (3)^2(0.02)] / [0.12 + 0.06 + 0.05 + 0.02]$$

$$= 1.76 \text{ and } \text{Var}[N_Q|N_P = 0] = E[N_Q^2|N_P = 0] - \{E[N_Q|N_P = 0]\}^2 = 1.76 - (0.88)^2$$

$$= 0.9856.$$

117. Solution: C

The domain of X and Y is pictured below. The shaded region is the portion of the domain over which $X < 0.2$.

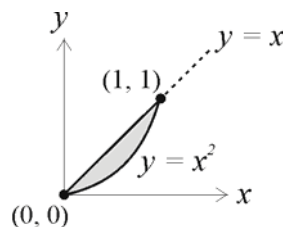


Now observe

$$\begin{aligned} \Pr[X < 0.2] &= \int_0^{0.2} \int_0^{1-x} 6[1 - (x + y)] dy dx = 6 \int_0^{0.2} \left[y - xy - \frac{1}{2}y^2 \right]_0^{1-x} dx \\ &= 6 \int_0^{0.2} \left[1 - x - x(1 - x) - \frac{1}{2}(1 - x)^2 \right] dx = 6 \int_0^{0.2} \left[(1 - x)^2 - \frac{1}{2}(1 - x)^2 \right] dx \\ &= 6 \int_0^{0.2} \frac{1}{2}(1 - x)^2 dx = -(1 - x)^3 \Big|_0^{0.2} = -(0.8)^3 + 1 = 0.488 \end{aligned}$$

118. Solution: E

The shaded portion of the graph below shows the region over which $f(x, y)$ is nonzero:



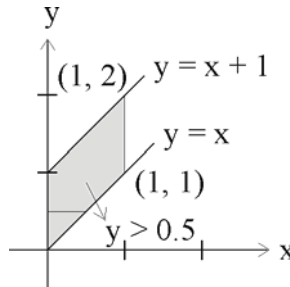
We can infer from the graph that the marginal density function of Y is given by

$$g(y) = \int_y^{\sqrt{y}} 15y dx = 15xy \Big|_y^{\sqrt{y}} = 15y(\sqrt{y} - y) = 15y^{3/2}(1 - y^{1/2}), \quad 0 < y < 1$$

or more precisely, $g(y) = \begin{cases} 15y^{3/2}(1-y)^{1/2}, & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$

119. Solution: D

The diagram below illustrates the domain of the joint density $f(x, y)$ of X and Y .



We are told that the marginal density function of X is $f_x(x) = 1$, $0 < x < 1$

while $f_{y|x}(y|x) = 1$, $x < y < x + 1$

It follows that $f(x, y) = f_x(x)f_{y|x}(y|x) = \begin{cases} 1 & \text{if } 0 < x < 1, x < y < x + 1 \\ 0 & \text{otherwise} \end{cases}$

Therefore,

$$\begin{aligned} \Pr[Y > 0.5] &= 1 - \Pr[Y \leq 0.5] = 1 - \int_0^{1/2} \int_x^{1/2} dy dx \\ &= 1 - \int_0^{1/2} y \Big|_x^{1/2} dx = 1 - \int_0^{1/2} \left(\frac{1}{2} - x \right) dx = 1 - \left[\frac{1}{2}x - \frac{1}{2}x^2 \right] \Big|_0^{1/2} = 1 - \frac{1}{4} + \frac{1}{8} = \frac{7}{8} \end{aligned}$$

[Note since the density is constant over the shaded parallelogram in the figure the solution is also obtained as the ratio of the area of the portion of the parallelogram above $y = 0.5$ to the entire shaded area.]

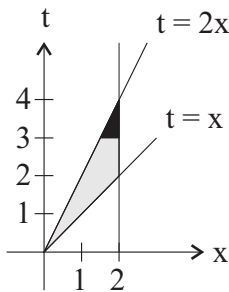
120. Solution: A

We are given that X denotes loss. In addition, denote the time required to process a claim by T .

Then the joint pdf of X and T is $f(x, t) = \begin{cases} \frac{3}{8}x^2 \cdot \frac{1}{x} = \frac{3}{8}x, & x < t < 2x, 0 \leq x \leq 2 \\ 0, & \text{otherwise.} \end{cases}$

Now we can find $P[T \geq 3] =$

$$\int_3^4 \int_{t/2}^2 \frac{3}{8} x dx dt = \int_3^4 \left[\frac{3}{16} x^2 \right]_{t/2}^2 dt = \int_3^4 \left(\frac{12}{16} - \frac{3}{64} t^2 \right) dt = \left[\frac{12}{16} t - \frac{1}{64} t^3 \right]_3^4 = \frac{12}{4} - 1 - \left(\frac{36}{16} - \frac{27}{64} \right) = 11/64 = 0.17.$$



121. Solution: C

The marginal density of X is given by

$$f_x(x) = \int_0^1 \frac{1}{64} (10 - xy^2) dy = \frac{1}{64} \left(10y - \frac{xy^3}{3} \right) \bigg|_0^1 = \frac{1}{64} \left(10 - \frac{x}{3} \right)$$

$$\begin{aligned} \text{Then } E(X) &= \int_2^{10} x f_x(x) dx = \int_2^{10} \frac{1}{64} \left(10x - \frac{x^2}{3} \right) dx = \frac{1}{64} \left(5x^2 - \frac{x^3}{9} \right) \bigg|_2^{10} \\ &= \frac{1}{64} \left[\left(500 - \frac{1000}{9} \right) - \left(20 - \frac{8}{9} \right) \right] = 5.778 \end{aligned}$$

122. Solution: D

The marginal distribution of Y is given by $f_2(y) = \int_0^y 6e^{-x} e^{-2y} dx = 6e^{-2y} \int_0^y e^{-x} dx$

$$= -6e^{-2y} e^{-x} + 6e^{-2y} = 6e^{-2y} - 6e^{-3y}, 0 < y < \infty$$

$$\text{Therefore, } E(Y) = \int_0^{\infty} y f_2(y) dy = \int_0^{\infty} (6ye^{-2y} - 6ye^{-3y}) dy = 6 \int_0^{\infty} ye^{-2y} dy - 6 \int_0^{\infty} ye^{-3y} dy =$$

$$\frac{6}{2} \int_0^{\infty} 2ye^{-2y} dy - \frac{6}{3} \int_0^{\infty} 3ye^{-3y} dy$$

But $\int_0^{\infty} 2ye^{-2y} dy$ and $\int_0^{\infty} 3ye^{-3y} dy$ are equivalent to the means of exponential random

variables with parameters 1/2 and 1/3, respectively. In other words, $\int_0^{\infty} 2ye^{-2y} dy = 1/2$

and $\int_0^{\infty} 3ye^{-3y} dy = 1/3$. We conclude that $E(Y) = (6/2)(1/2) - (6/3)(1/3) = 3/2 - 2/3 = 9/6 - 4/6 = 5/6 = 0.83$.

123. Solution: C

Observe

$$\begin{aligned} \Pr[4 < S < 8] &= \Pr[4 < S < 8 | N = 1] \Pr[N = 1] + \Pr[4 < S < 8 | N > 1] \Pr[N > 1] \\ &= \frac{1}{3} \left(e^{-4/5} - e^{-8/5} \right) + \frac{1}{6} \left(e^{-1/2} - e^{-1} \right) * \\ &= 0.122 \end{aligned}$$

*Uses that if X has an exponential distribution with mean μ

$$\Pr(a \leq X \leq b) = \Pr(X \geq a) - \Pr(X \geq b) = \int_a^{\infty} \frac{1}{\mu} e^{-t/\mu} dt - \int_b^{\infty} \frac{1}{\mu} e^{-t/\mu} dt = e^{-\frac{a}{\mu}} - e^{-\frac{b}{\mu}}$$

124. Solution: A

First note that $f_{(X,Y)}(x,y) = e^{-x} \cdot (2e^{-2y}) = f_X(x) \cdot f_Y(y)$ and that the support of $f_{(X,Y)}$ is a cross product. Therefore X and Y are independent. Thus

$$\begin{aligned} & \text{Var}(Y | X > 3 \text{ and } Y > 3) \\ &= \text{Var}(Y | Y > 3) \\ &= \text{Var}(Y + 3) \\ &= \text{Var}(Y) + \text{Var}(3) \\ &= \text{Var}(Y) \\ &= E(Y)^2 \\ &= (0.5)^2 \\ &= 0.25 \end{aligned}$$

Independence of X and Y
Memory-less property of exponential
Independence of Y and 3
 $\text{Var}(3) = 0$
Exponential variance

125. Solution: E

The support of (X,Y) is $0 < y < x < 1$.

$f_{X,Y}(x,y) = f(y|x)f_X(x) = 2$ on that support. It is clear geometrically

(a flat joint density over the triangular region $0 < y < x < 1$) that when $Y = y$

we have $X \sim U(y, 1)$ so that $f(x|y) = \frac{1}{1-y}$ for $y < x < 1$.

By computation:

$$f_Y(y) = \int_y^1 2dx = 2 - 2y \Rightarrow f(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{2}{2-2y} = \frac{1}{1-y} \text{ for } y < x < 1$$

126. Solution: C

Using the notation of the problem, we know that $p_0 + p_1 = \frac{2}{5}$ and

$$p_0 + p_1 + p_2 + p_3 + p_4 + p_5 = 1.$$

Let $p_n - p_{n+1} = c$ for all $n \leq 4$. Then $p_n = p_0 - nc$ for $1 \leq n \leq 5$.

$$\text{Thus } p_0 + (p_0 - c) + (p_0 - 2c) + \dots + (p_0 - 5c) = 6p_0 - 15c = 1.$$

$$\text{Also } p_0 + p_1 = p_0 + (p_0 - c) = 2p_0 - c = \frac{2}{5}. \text{ Solving simultaneously } \begin{cases} 6p_0 - 15c = 1 \\ 2p_0 - c = \frac{2}{5} \end{cases}$$

$$\begin{aligned} 6p_0 - 3c &= \frac{6}{5} \\ \Rightarrow \frac{-6p_0 + 15c}{12c} &= \frac{-1}{5}. \text{ So } c = \frac{1}{60} \text{ and } 2p_0 = \frac{2}{5} + \frac{1}{60} = \frac{25}{60}. \text{ Thus } p_0 = \frac{25}{120}. \end{aligned}$$

$$\text{We want } p_4 + p_5 = (p_0 - 4c) + (p_0 - 5c) = \frac{17}{120} + \frac{15}{120} = \frac{32}{120} = 0.267.$$

127. Solution: D

Because the number of payouts (including payouts of zero when the loss is below the deductible) is large, we can apply the central limit theorem and assume the total payout S is normal. For one loss there is no payout with probability 0.25 and otherwise the payout is $U(0, 15000)$. So,

$$E[X] = 0.25 * 0 + 0.75 * 7500 = 5625,$$

$$E[X^2] = 0.25 * 0 + 0.75 * (7500^2 + \frac{15000^2}{12}) = 56,250,000, \text{ so the variance of one claim is}$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = 24,609,375.$$

Applying the CLT,

$$P[1,000,000 < S < 1,200,000] = P\left[-1.781741613 < \frac{S - (200)(5625)}{\sqrt{(200)(24,609,375)}} < 1.069044968\right]$$

which interpolates to $0.8575 - (1 - 0.9626) = 0.8201$ from the provided table.

128. Key: B

Let H be the percentage of clients with homeowners insurance and R be the percentage of clients with renters insurance.

Because 36% of clients do not have auto insurance and none have both homeowners and renters insurance, we calculate that 8% ($36\% - 17\% - 11\%$) must have renters insurance, but not auto insurance.

$(H - 11)\%$ have both homeowners and auto insurance, $(R - 8)\%$ have both renters and auto insurance, and none have both homeowners and renters insurance, so $(H + R - 19)\%$ must equal 35%. Because $H = 2R$, R must be 18%, which implies that 10% have both renters and auto insurance.

129. Key: B

The reimbursement is positive if health care costs are greater than 20, and because of the memoryless property of the exponential distribution, the conditional distribution of health care costs greater than 20 is the same as the unconditional distribution of health care costs.

We observe that a reimbursement of 115 corresponds to health care costs of 150 ($100\% \times (120 - 20) + 50\% \times (150 - 120)$), which is 130 greater than the deductible of 20.

Therefore, $G(115) = F(130) = 1 - e^{-\frac{130}{100}} = 0.727$.

130. Key: C

$$E[100(0.5)^X] = 100E[(0.5)^X] = 100E[e^{(\ln 0.5)X}] = 100M_X(\ln 0.5) = 100 \frac{1}{1 - 2 \ln 0.5} = 41.9$$

131. Solution: E

First, find the conditional probability function of N_2 given $N_1 = n_1$: $p_{2|1}(n_2 | n_1) = \frac{p(n_1, n_2)}{p_1(n_1)}$,

where $p_1(n_1)$ is the marginal probability function of N_1 . To find the latter, sum the joint probability function over all possible values of N_2 obtaining

$$p_1(n_1) = \sum_{n_2=1}^{\infty} p(n_1, n_2) = \frac{3}{4} \left(\frac{1}{4} \right)^{n_1-1} \sum_{n_2=1}^{\infty} e^{-n_1} (1 - e^{-n_1})^{n_2-1} = \frac{3}{4} \left(\frac{1}{4} \right)^{n_1-1},$$

since $\sum_{n_2=1}^{\infty} e^{-n_1} (1 - e^{-n_1})^{n_2-1} = 1$ as the sum of the probabilities of a geometric random variable. The conditional probability function is

$$p_{2|1}(n_2 | n_1) = \frac{p(n_1, n_2)}{p_1(n_1)} = e^{-n_1} (1 - e^{-n_1})^{n_2-1},$$

which is the probability function of a geometric random variable with parameter $p = e^{-n_1}$. The mean of this distribution is $1/p = 1/e^{-n_1} = e^{n_1}$, and becomes e^2 when $n_1 = 2$.

132. Solution: C

The number of defective modems is $20\% \times 30 + 8\% \times 50 = 10$.

The probability that exactly two of a random sample of five are defective is $\frac{\binom{10}{2} \binom{70}{3}}{\binom{80}{5}} = 0.102$.

133. Solution: B

$$\begin{aligned} \Pr(\text{man dies before age 50}) &= \Pr(T < 50 | T > 40) \\ &= \frac{\Pr(40 < T < 50)}{\Pr(T > 40)} = \frac{F(50) - F(40)}{1 - F(40)} \\ &= \frac{e^{\frac{1-1.1^{40}}{1000}} - e^{\frac{1-1.1^{50}}{1000}}}{e^{\frac{1-1.1^{40}}{1000}}} = 1 - e^{\frac{(1.1^{40} - 1.1^{50})}{1000}} \\ &= 0.0696 \end{aligned}$$

$$\text{Expected Benefit} = 5000 \Pr(\text{man dies before age 50}) = (5000) (0.0696) = 347.96$$

134. Solutions: C

Letting t denote the relative frequency with which twin-sized mattresses are sold, we have that the relative frequency with which king-sized mattresses are sold is $3t$ and the relative frequency with which queen-sized mattresses are sold is $(3t+t)/4$, or t . Thus, $t = 0.2$ since $t + 3t + t = 1$. The probability we seek is $3t + t = 0.80$.

135. Key: E

$$\text{Var}(N) = E[\text{Var}(N|\lambda)] + \text{Var}[E(N|\lambda)] = E(\lambda) + \text{Var}(\lambda) = 1.50 + 0.75 = 2.25$$

136. Key: D

X follows a geometric distribution with $p = \frac{1}{6}$. $Y = 2$ implies the first roll is not a 6 and the second roll is a 6. This means a 5 is obtained for the first time on the first roll (probability = 20%) or a 5 is obtained for the first time on the third or later roll (probability = 80%).

$$E[X | X \geq 3] = \frac{1}{p} + 2 = 6 + 2 = 8, \text{ so } E[X|Y = 2] = 0.2(1) + 0.8(8) = 6.6$$

137. Key: E

Because X and Y are independent and identically distributed, the moment generating function of $X + Y$ equals $K^2(t)$, where $K(t)$ is the moment generating function common to X and Y . Thus, $K(t) = 0.30e^{-t} + 0.40 + 0.30e^t$. This is the moment generating function of a discrete random variable that assumes the values -1, 0, and 1 with respective probabilities 0.30, 0.40, and 0.30. The value we seek is thus 0.70.

138. Key: D

Suppose the component represented by the random variable X fails last. This is represented by the triangle with vertices at $(0, 0)$, $(10, 0)$ and $(5, 5)$. Because the density is uniform over this region, the mean value of X and thus the expected operational time of the machine is 5. By symmetry, if the component represented by the random variable Y fails last, the expected operational time of the machine is also 5. Thus, the unconditional expected operational time of the machine must be 5 as well.

139. Key: B

The unconditional probabilities for the number of people in the car who are hospitalized are 0.49, 0.42 and 0.09 for 0, 1 and 2, respectively. If the number of people hospitalized is 0 or 1, then the total loss will be less than 1. However, if two people are hospitalized, the probability that the total loss will be less than 1 is 0.5. Thus, the expected number of people in the car who are hospitalized, given that the total loss due to hospitalizations from the accident is less than 1 is

$$\frac{0.49}{0.49 + 0.42 + 0.09 \cdot 0.5} \cdot 0 + \frac{0.42}{0.49 + 0.42 + 0.09 \cdot 0.5} \cdot 1 + \frac{0.09 \cdot 0.5}{0.49 + 0.42 + 0.09 \cdot 0.5} \cdot 2 = 0.534$$

140. Key: B

Let X equal the number of hurricanes it takes for two losses to occur. Then X is negative binomial with “success” probability $p = 0.4$ and $r = 2$ “successes” needed.

$$P[X = n] = \binom{n-1}{r-1} p^r (1-p)^{n-r} = \binom{n-1}{2-1} (0.4)^2 (1-0.4)^{n-2} = (n-1)(0.4)^2 (0.6)^{n-2}, \text{ for } n \geq 2.$$

We need to maximize $P[X = n]$. Note that the ratio

$$\frac{P[X = n+1]}{P[X = n]} = \frac{n(0.4)^2 (0.6)^{n-1}}{(n-1)(0.4)^2 (0.6)^{n-2}} = \frac{n}{n-1} (0.6).$$

This ratio of “consecutive” probabilities is greater than 1 when $n = 2$ and less than 1 when $n \geq 3$. Thus, $P[X = n]$ is maximized at $n = 3$; the mode is 3.

141. Key: C

There are 10 ($5 \text{ choose } 3$) ways to select the three columns in which the three items will appear. The row of the rightmost selected item can be chosen in any of six ways, the row of the leftmost selected item can then be chosen in any of five ways, and the row of the middle selected item can then be chosen in any of four ways. The answer is thus $(10)(6)(5)(4) = 1200$. Alternatively, there are 30 ways to select the first item. Because there are 10 squares in the row or column of the first selected item, there are $30 - 10 = 20$ ways to select the second item. Because there are 18 squares in the rows or columns of the first and second selected items, there are $30 - 18 = 12$ ways to select the third item. The number of permutations of three qualifying items is $(30)(20)(12)$. The number of combinations is thus $(30)(20)(12)/3! = 1200$.

142. Key: B

The expected bonus for a high-risk driver is $0.8 \cdot 12 \text{ (months)} \cdot 5.00 = 48$.

The expected bonus for a low-risk driver is $0.9 \cdot 12 \text{ (months)} \cdot 5.00 = 54$.

The expected bonus payment from the insurer is $600 \cdot 48 + 400 \cdot 54 = 50,400$.

143. Key: E

$P(\text{Pr} \cup \text{Li}) = P(\text{Pr}) + P(\text{Li} \cap \text{Pr}') = 0.10 + 0.01$. Subtract from 1 to get the answer.

144. Key: E

The total time is less than 60 minutes, so if x minutes are spent in the waiting room, less than $60 - x$ minutes are spent in the meeting itself.

145. Key: C

$$f(y | x = 0.75) = \frac{f(0.75, y)}{\int_0^1 f(0.75, y) dy} = \frac{f(0.75, y)}{1.125}.$$

Thus,

$$f(y | x = 0.75) = \begin{cases} 4/3 & \text{for } 0 < y < 0.5 \\ 2/3 & \text{for } 0.5 < y < 1 \end{cases},$$

which leads to $\text{Var}(Y | X = 0.75) = 11/144 = 0.076$.

146. Key: B

C = the set of TV watchers who watched CBS over the last year

N = the set of TV watchers who watched NBC over the last year

A = the set of TV watchers who watched ABC over the last year

H = the set of TV watchers who watched HGTV over the last year

The number of TV watchers in the set $C \cup N \cup A$ is $34 + 15 + 10 - 7 - 6 - 5 + 4 = 45$.

Because $C \cup N \cup A$ and H are mutually exclusive, the number of TV watchers in the set $C \cup N \cup A \cup H$ is $45 + 18 = 63$.

The number of TV watchers in the complement of $C \cup N \cup A \cup H$ is thus $100 - 63 = 37$.

147. Key: A

Let X denote the amount of a claim before application of the deductible. Let Y denote the amount of a claim payment after application of the deductible. Let λ be the mean of X , which because X is exponential, implies that λ^2 is the variance of X and $E(X^2) = 2\lambda^2$.

By the memoryless property of the exponential distribution, the conditional distribution of the portion of a claim above the deductible given that the claim exceeds the deductible is an exponential distribution with mean λ . Given that $E(Y) = 0.9\lambda$, this implies that the probability of a claim exceeding the deductible is 0.9 and thus $E(Y^2) = 0.9 \cdot 2\lambda^2 = 1.8\lambda^2$. Then,
 $\text{Var}(Y) = 1.8\lambda^2 - (0.9\lambda)^2 = 0.99\lambda^2$.

148. Key: C

Let N denote the number of hurricanes, which is Poisson distributed with mean and variance 4.

Let X_i denote the loss due to the i^{th} hurricane, which is exponentially distributed with mean 1,000 and therefore variance $(1,000)^2 = 1,000,000$.

Let X denote the total loss due to the N hurricanes.

This problem can be solved using the conditional variance formula. Note that independence is used to write the variance of a sum as the sum of the variances.

$$\begin{aligned}\text{Var}(X) &= \text{Var}[E(X|N)] + E[\text{Var}(X|N)] \\ &= \text{Var}[E(X_1 + \dots + X_N)] + E[\text{Var}(X_1 + \dots + X_N)] \\ &= \text{Var}[NE(X_1)] + E[N\text{Var}(X_1)] \\ &= \text{Var}(1,000N) + E(1,000,000N) \\ &= 1,000^2 \text{Var}(N) + 1,000,000E(N) \\ &= 1,000,000(4) + 1,000,000(4) = 8,000,000\end{aligned}$$

149. Key: B

Let N denote the number of accidents, which is binomial with parameters $\frac{1}{4}$ and 3 and thus has mean $3\left(\frac{1}{4}\right) = \frac{3}{4}$ and variance $3\left(\frac{1}{4}\right)\left(\frac{3}{4}\right) = \frac{9}{16}$.

Let X_i denote the unreimbursed loss due to the i^{th} accident, which is 0.3 times an exponentially distributed random variable with mean 0.8 and therefore variance $(0.8)^2 = 0.64$. Thus, X_i has mean $0.8(0.3) = 0.24$ and variance $0.64(0.3)^2 = 0.0576$.

Let X denote the total unreimbursed loss due to the N accidents.

This problem can be solved using the conditional variance formula. Note that independence is used to write the variance of a sum as the sum of the variances.

$$\begin{aligned}\text{Var}(X) &= \text{Var}\left[E(X|N)\right] + E\left[\text{Var}(X|N)\right] \\ &= \text{Var}\left[E(X_1 + \dots + X_N)\right] + E\left[\text{Var}(X_1 + \dots + X_N)\right] \\ &= \text{Var}\left[NE(X_1)\right] + E\left[N\text{Var}(X_1)\right] \\ &= \text{Var}(0.24N) + E(0.0576N) \\ &= (0.24)^2 \text{Var}(N) + 0.0576E(N) \\ &= 0.0576\left(\frac{9}{16}\right) + 0.0576\left(\frac{3}{4}\right) = 0.0756.\end{aligned}$$

150. Key: B

The 95th percentile is in the range when an accident occurs. It is the 75th percentile of the payout, given that an accident occurs, because $(0.95 - 0.80)/(1 - 0.80) = 0.75$. Letting x be the 75th percentile of the given exponential distribution, $F(x) = 1 - e^{-\frac{x}{3000}} = 0.75$, so $x = 4159$. Subtracting the deductible of 500 gives 3659 as the (unconditional) 95th percentile of the insurance company payout.

151. Key: C

The ratio of the probability that one of the damaged pieces is insured to the probability

$$\text{that none of the damaged pieces are insured is } \frac{\frac{\binom{r}{1}\binom{27-r}{3}}{\binom{27}{4}}}{\frac{\binom{r}{0}\binom{27-r}{4}}{\binom{27}{4}}} = \frac{4r}{24-r}, \text{ where } r \text{ is the total}$$

number of pieces insured. Setting this ratio equal to 2 and solving yields $r = 8$.

The probability that two of the damaged pieces are insured is

$$\frac{\frac{\binom{r}{2}\binom{27-r}{2}}{\binom{27}{4}}}{\frac{\binom{8}{2}\binom{19}{2}}{\binom{27}{4}}} = \frac{(8)(7)(19)(18)(4)(3)(2)(1)}{(27)(26)(25)(24)(2)(1)(2)(1)} = \frac{266}{975} = 0.27.$$

152. Key: A

The probability that Rahul examines exactly n policies is $0.1 \cdot 0.9^{n-1}$. The probability that Toby examines more than n policies is 0.8^n . The required probability is thus

$$\sum_{n=1}^{\infty} 0.1 \cdot 0.9^{n-1} \cdot 0.8^n = \frac{1}{9} \sum_{n=1}^{\infty} 0.72^n = \frac{0.72}{9 \cdot 1 - 0.72} = 0.2857.$$

An alternative solution begins by imagining Rahul and Toby examine policies simultaneously until at least one of the finds a claim. At each examination there are four possible outcomes:

1. Both find a claim. The probability is 0.02.
2. Rahul finds a claim and Toby does not. The probability is 0.08.
3. Toby finds a claim and Rahul does not. The probability is 0.18
4. Neither finds a claim. The probability is 0.72.

Conditioning on the examination at which the process ends, the probability that it ends with Rahul being the first to find a claim (and hence needing to examine fewer policies) is $0.08/(0.02 + 0.08 + 0.18) = 8/28 = 0.2857$.

153. KEY: E

Solution:

1. For a Poisson random variable N , $E(N) = \text{Var}(N) = \lambda$ and $E(N^2) = \lambda + \lambda^2$.
2. We are given $\lambda_x = \lambda_y - 8$ and $\lambda_x + \lambda_x^2 = 0.6 * (\lambda_y + \lambda_y^2)$.
3. Combining equations, we get a quadratic in λ_y : $\lambda_y - 8 + (\lambda_y - 8)^2 = 0.6 * (\lambda_y + \lambda_y^2)$.
4. Simplifying, we get: $0.4 \lambda_y^2 - 15.6 \lambda_y + 56 = 0$.
5. Two roots: $\lambda_y = 35$ and 4 . The second root is rejected since this would make λ_x negative.
6. Therefore, the variance is 35 .